

Discrete random walks (1)

A-First results (obtained "by hand", no formalism)

1) Return probability: R .

a) Bayes theorem: $R = pR_{\rightarrow} + qR_{\leftarrow}$

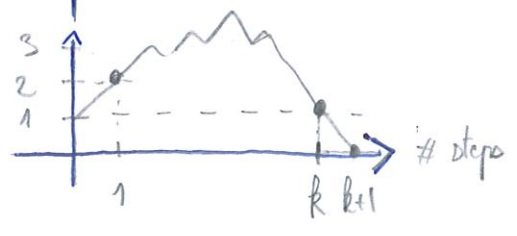
b) Consider a walker that starts from site 1; it can visit site 0 after n visits to site 1, with $n = 1, 2, \dots, \infty \rightarrow$ the different values of n form a partition of all possible events:

$n=1$



Prob q for this event

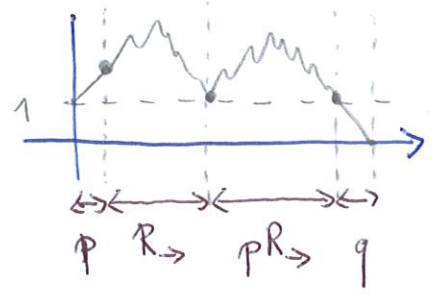
$n=2$



walker starts at site 1, moves right, (prob p), then goes back to 1 (prob R_{\rightarrow}) and from site 1, moves left to 0 (prob q)

\hookrightarrow prob $pR_{\rightarrow}q$

$n=3$



Prob: $(pR_{\rightarrow})^2 q$

...

$$\Rightarrow R_{\rightarrow} = q \sum_{n=1}^{\infty} (pR_{\rightarrow})^{n-1} = \frac{q}{1 - pR_{\rightarrow}}$$

$$\Leftrightarrow pR_{\rightarrow}^2 - R_{\rightarrow} + q = 0$$

$$\Delta = 1 - 4pq = 1 - 4p(1-p) = 4p^2 - 4p + 1 = (2p-1)^2 = (p-q)^2$$

$$R_{\rightarrow} = \frac{1}{2p} (1 \pm |p-q|)$$

Root with \oplus $\xrightarrow{p > 1/2}$ $\frac{1}{2p} (1 + p - q) = \frac{1}{2p} (p + p) = 1$: makes no sense (take $p \rightarrow 1$, walker should not come back)
 $\xrightarrow{p < 1/2}$ $\frac{1}{2p} (1 - p + q) = \frac{q}{p} > 1$: no!

This root makes no sense.

Look at the other root:

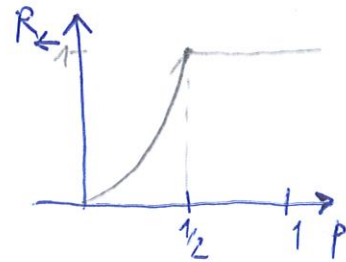
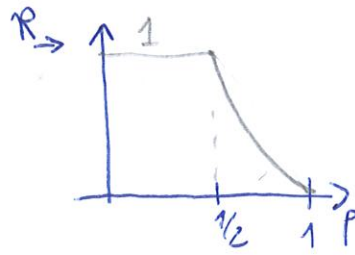
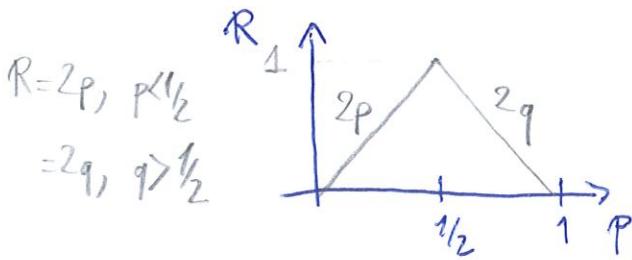
$$R_{\rightarrow} = \frac{1}{2p} (1 - p + q) = q/p, \quad p \geq \frac{1}{2} \quad \text{ok}$$

$$= \frac{1}{2p} (1 + q + p) = 1, \quad p \leq \frac{1}{2} \quad \text{ok}$$

Thus $R_{\rightarrow} = \frac{1}{2p} (1 - |p - q|)$; $R_{\leftarrow} = \frac{1}{2q} (1 - |p - q|)$, permuting p & q

c) $R = p R_{\rightarrow} + q R_{\leftarrow}$;

$R = 1 - |p - q|$; For $p = q = 1/2$, $R = 1$
return is certain.



We have found, for instance with $p \geq 1/2$: $R_{\rightarrow} = q/p$.

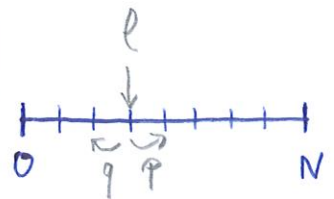
This also tells us that starting from site $l > 0$, the proba to visit the origin, some time in the future, is simply $R_{l \rightarrow} = (q/p)^l$.

On the other hand, with $p \leq 1/2$, $R_{l \rightarrow} = 1^l = 1$.

2) Exit time from an interval

a) $P[T(l) = m+1] = p P[T(l+1) = m] + q P[T(l-1) = m]$

$P[T(0) = n] = \delta_{n,0}$; $P[T(N) = n] = \delta_{n,0}$



b) $\langle T(l) \rangle = \sum_{n=0}^{\infty} n P[T(l) = n] = \sum_{n=1}^{\infty} n P[T(l) = n]$

$$= \sum_{n=1}^{\infty} n \left\{ p P[T(l+1) = n] + q P[T(l-1) = n-1] \right\}$$

$$= \sum_{n=0}^{\infty} (n+1) \left\{ p P[T(l+1) = n] + q P[T(l-1) = n] \right\}$$

$$= \sum_{n=0}^{\infty} n \left\{ p P[T(l+1) = n] + q P[T(l-1) = n] \right\} + \sum_{n=0}^{\infty} 1 \cdot \{ p P + q P \}$$

$$\boxed{\langle T(l) \rangle = p \langle T(l+1) \rangle + q \langle T(l-1) \rangle + \frac{p+q}{1}}; \quad \langle T(0) \rangle = 0$$

$$\langle T(N) \rangle = 0$$

Discrete random walks (2)

c) Resolution of the difference equation $p x_{l+1} + q x_{l-1} - x_l = -1$; $x_0 = x_N = 0$

For $p=q=1/2$, we recognize a discrete Laplacian.

→ homogeneous equation : $p x_{l+1} + q x_{l-1} - x_l = 0$; take $x_l = r^l$
 $p r^2 - r + q = 0$; already solved above.

we have found the two roots to be 1 and q/p ; we assume here $q \neq p$
 The general solution is thus of form $\lambda + \mu (q/p)^l$

→ we need a particular solution, that we can take polynomial.

A constant does not work since $p+q-1=0$, but a linear function works

$$x_l = b l \Rightarrow p b(l+1) + q b(l-1) - b l = -1 \Leftrightarrow p b - q b = -1$$

$$b = \frac{1}{q-p}$$

→ the complete solution reads:

$$x_l = \lambda + \mu \left(\frac{q}{p}\right)^l + \frac{l}{q-p}$$

Boundary conditions : $x_0 = x_N = 0 \Rightarrow \lambda + \mu = 0$; $\lambda + \mu \left(\frac{q}{p}\right)^N + \frac{N}{q-p} = 0$

$$\lambda \left[1 - \left(\frac{q}{p}\right)^N \right] = -\frac{N}{q-p}$$

Thus: $\langle T(l) \rangle = \frac{1}{q-p} \left[l - N \frac{1 - (q/p)^l}{1 - (q/p)^N} \right]$ for $q \neq p$.
 (ie $p \neq \frac{1}{2}$)

→ Particular case $p=q=1/2$: the 2 roots above, 1 and q/p , become equal.

We can no longer look for a solution in $\alpha \pi_1^l + \beta \pi_2^l$, but in $\alpha \pi_1^l + \beta l \pi_2^l$. Here, $\pi_1 = \pi_2 = 1 \rightarrow$ the general solution to the homogeneous equation is in $\alpha + \beta l$.

The particular solution can be taken in l^2

It has to be symmetric wrt $l = \frac{N}{2} \rightarrow \left[\left(l - \frac{N}{2} \right)^2 + \text{Cst} \right] \times \gamma$

The constant is set by the b.c :

$$\hookrightarrow \gamma \left\{ \left(l - \frac{N}{2} \right)^2 - \left(\frac{N}{2} \right)^2 \right\} = \gamma l(N-l), \text{ to be plugged into}$$

$$\frac{1}{2} x_{l+1} - x_l + \frac{1}{2} x_{l-1} = -1$$

$$\Rightarrow \langle T(l) \rangle = l(N-l)$$

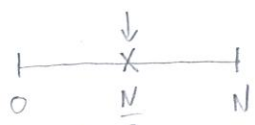
Alternatively, one may take the expression of $\langle T(l) \rangle$ for $p \neq q$, and take the limit $p \rightarrow \frac{1}{2}$ (i.e. $q \rightarrow \frac{1}{2}$ also); gives same result

Remark 1 it is possible to show that the walker leaves $[0, N]$ with proba one, when waiting enough. Indeed, denote $\pi(l)$ the proba to leave the interval (from its right or its left end), one day.

$$\pi(l) = p \pi(l+1) + q \pi(l-1) \quad \text{with} \quad \pi(0) = \pi(N) = 1$$

$$\Rightarrow \pi(l) = \alpha + \beta \left(\frac{q}{p}\right)^l \quad ; \quad \begin{cases} \alpha + \beta = 1 \\ \alpha + \beta \left(\frac{q}{p}\right)^N = 1 \end{cases} \Rightarrow \begin{cases} \alpha = 1 \\ \beta = 0 \end{cases}$$

$$\Rightarrow \pi(l) = 1$$

Remark 2 : take $l = \frac{N}{2}$  ; $\langle T(\frac{N}{2}) \rangle = \frac{N^2}{4} \propto N^2$, diffusive scaling

Here, the diffusion coef is $D = \frac{1}{2}$

$$\hookrightarrow \text{position} = \sum_{i=1}^m +1 \alpha_i - 1$$

$$v(\text{position}) = m = 2 \times \underset{\downarrow \frac{1}{2}}{D} \times \underset{\downarrow \text{time}}{m}$$

$$\text{time} \propto \frac{\text{distance}^2}{D} \Rightarrow \text{diff coefficient}$$

Discrete random walks (3)

(B) Generating function formalism

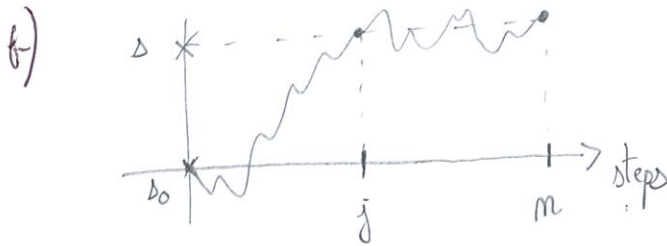
1) a) $F_n(\Delta|D_0) \leq P_n(\Delta|D_0)$

$\sum_{\Delta} P_n(\Delta|D_0) = 1$ the walker has to be somewhere

$\sum_n P_n(\Delta|D_0)$ may diverge (mean # of visits, see below)

$\sum_n F_n(\Delta|D_0) = R(\Delta|D_0) \leq 1$: proba to reach site Δ , some time, starting at D_0

$\sum_{\Delta} F_{\Delta}(\Delta|D_0) \leq 1$



When the walker is at Δ , it has first visited this site earlier than step n , at step j with $j = 1, 2, \dots, n$ (for $n \geq 1$) and starting from Δ at step j it has to be back at step n

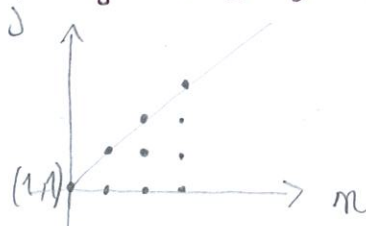
For $n \geq 1$: $P_n(\Delta|D_0) = \sum_{j=1}^n F_j(\Delta|D_0) P_{n-j}(\Delta|\Delta)$

Law of total probability
 $P_n(\Delta|D_0) = \sum_j \text{Prob}(\text{beats } \Delta | j) P_j$

For $n = 0$: $P_0(\Delta|D_0) = \delta_{\Delta, D_0}$

i.e. $P_n(\Delta|D_0) = \delta_{n,0} \delta_{\Delta, D_0} + \sum_{j=1}^n F_j(\Delta|D_0) P_{n-j}(\Delta|\Delta)$
 interpreting the \sum_1^0 to be zero (for $n = 0$).

c) $P(\Delta|D_0; \xi) = \sum_{n=0}^{\infty} P_n(\Delta|D_0) \xi^n = \delta_{\Delta, D_0} + \sum_{n=1}^{\infty} P_n(\Delta|D_0) \xi^n$
 $= \delta_{\Delta, D_0} + \sum_{n=1}^{\infty} \sum_{j=1}^n F_j(\Delta|D_0) P_{n-j}(\Delta|\Delta) \xi^j \xi^{n-j}$



$1 \leq n \leq \infty$
 $1 \leq j \leq n$

$= \delta_{\Delta, D_0} + \sum_{j=1}^{\infty} \sum_{k=0}^{\infty} F_j(\Delta|D_0) \xi^j P_k(\Delta|\Delta) \xi^k$
 $= \sum_{j=0}^{\infty} \text{since } F_0 = 0$

$\Leftrightarrow \begin{cases} 1 \leq j \leq \infty \\ n-j \geq 0 \\ k \end{cases}$

$P(\Delta|D_0; \xi) = \delta_{\Delta, D_0} + F(\Delta|D_0; \xi) P(D_0|D_0; \xi)$

2) We have found that $P(\Delta|\Delta; \xi) = 1 + F(\Delta|\Delta; \xi)P(\Delta|\Delta; \xi)$

and we want that $R(\Delta|\Delta) = 1$. Note that

$$R(\Delta|\Delta) = \sum_{n=0}^{\infty} F_n(\Delta|\Delta) = \lim_{\xi \rightarrow 1} F(\Delta|\Delta; \xi)$$

$$F(\Delta|\Delta; \xi) = 1 - \frac{1}{P(\Delta|\Delta; \xi)}$$

$$P(\Delta|\Delta; \xi) = \frac{1}{1 - F(\Delta|\Delta; \xi)}$$

$$R(\Delta|\Delta) = 1 \Leftrightarrow P(\Delta|\Delta; 1) = \infty$$

$$R(\Delta|\Delta) = F(\Delta|\Delta; 1)$$

useful for part C.

3) $I_n(\Delta|\Delta_0) = 1$ if site Δ occupied at time n ; 0 otherwise; $\langle I_n(\Delta|\Delta_0) \rangle = P_m(\Delta|\Delta_0)$

$\Rightarrow \sum_{n=0}^{\infty} I_n(\Delta|\Delta_0)$ is the total number of visits to site Δ .

$$\Rightarrow \left\langle \sum_{n=0}^{\infty} I_n(\Delta|\Delta_0) \right\rangle = \sum_{n=0}^{\infty} P_m(\Delta|\Delta_0) = P(\Delta|\Delta_0; 1) \quad \text{mean number of visits}$$

the mean number of returns is $P(\Delta|\Delta_0; 1) - 1 = \infty$ for recurrent walk

$$= \frac{1}{1 - R(\Delta|\Delta_0)} - 1 \quad \text{for transient } \Delta \text{ (non recurrent)}$$

Example: take $R = 0,99$;

Views each return like a new experiment, where walker comes back 99 times in 100 trials, and escapes to ∞ since.

$$\langle \# \text{ returns} \rangle = \frac{1}{0,01} - 1 = 99$$

More generally, for R given, $1-R$ is proba for not returning = $\frac{1}{\langle \# \text{ return} \rangle + 1}$

$$\Rightarrow \langle \# \text{ return} \rangle = \frac{1}{1-R} - 1$$

↑
1 trial not returning

4) $F_m(\Delta|\Delta_0)$ is the proba to visit Δ for the first time at step m .

Since $R(\Delta|\Delta_0) = \sum_{n=0}^{\infty} F_n(\Delta|\Delta_0)$, $\frac{F_m(\Delta|\Delta_0)}{R(\Delta|\Delta_0)}$ is the conditional proba to visit Δ for the first time at step m , given that site Δ is indeed visited.

$$\Rightarrow \left\langle Z(\Delta|\Delta_0) \right\rangle = \sum_{m=0}^{\infty} m \frac{F_m(\Delta|\Delta_0)}{R(\Delta|\Delta_0)} = \frac{F'(\Delta|\Delta_0; 1)}{F(\Delta|\Delta_0; 1)}$$

Discrete random walks (4)

(B) 5c) Application to biased r.w

$$P(D_0|D_0; \xi) = \sum_{n=0}^{\infty} P_n(D_0|D_0) \xi^n$$

$$= \sum_{i=0}^{\infty} P_{2i}(D_0|D_0) \xi^{2i}$$

$$= \sum_{i=0}^{\infty} \binom{2i}{i} p^i q^i \xi^{2i}$$

$$P(D_0|D_0; \xi) = (1 - 4pq\xi^2)^{-1/2}$$

$$; P_m(D_0|D_0) = 0 \text{ if } m \text{ is odd}$$

$$= \binom{n}{n/2} p^{n/2} q^{n/2} \text{ if } n \text{ even}$$

$$F(D_0|D_0; \xi) = 1 - \frac{1}{P(D_0|D_0; \xi)} = 1 - (1 - 4pq\xi^2)^{1/2}$$

$$\Rightarrow R(D_0|D_0) = F(D_0|D_0; 1) = 1 - (1 - 4pq)^{1/2} = 1 - |p - q|$$

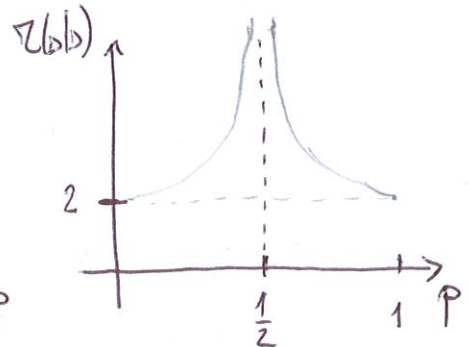
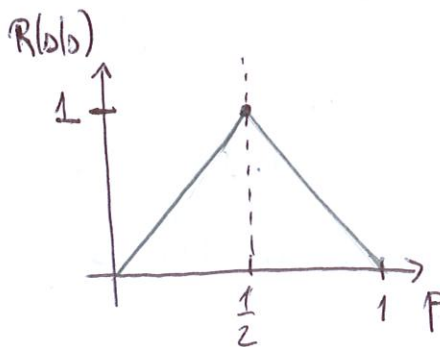
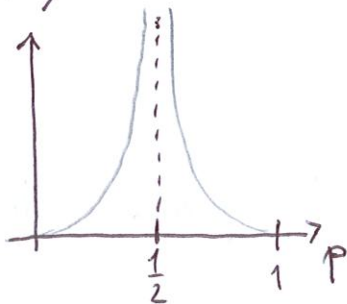
We recover the result of (A)

Mean # of returns = $P(D_0|D_0; 1) - 1 = \frac{1}{|p - q|} - 1$

Conditional mean first passage time:

$$Z(D_0|D_0) = \frac{F'(D_0|D_0; 1)}{F(D_0|D_0; 1)} = \frac{1}{1 - |p - q|} \cdot \frac{1}{2} \left(\frac{4pq \cdot 2\xi}{\sqrt{1 - 4pq\xi^2}} \right)_{\xi=1} = \frac{4pq}{(1 - |p - q|) |p - q|}$$

<# returns>



Note that $Z(b|b) \searrow$ when $p \nearrow$ and $\rightarrow 1$: the return probability decreases, which does not affect $Z(b|b)$, conditioned onto those walks that return.

When $p \rightarrow 1$, the only event leading to a return is a sequence of 2 steps, left/right. The only "possible" returns are in 2 steps.

Note that $Z(b|b) \rightarrow \infty$ for $p \rightarrow \frac{1}{2}$: although the walker always comes back to its starting point, it takes ∞ time on average.

c) translationally invariant random walks

1) a) $P_{m+1}(\vec{l}) = \sum_{\vec{l}'} P_m(\vec{l} - \vec{l}') p(\vec{l}')$

b) $\tilde{P}_{m+1}(\vec{k}) = \tilde{P}_m(\vec{k}) \lambda(\vec{k}) ; \tilde{P}_0(\vec{k}) = 1 \Rightarrow \tilde{P}_m(\vec{k}) = \lambda^m(\vec{k})$

$P_m(\vec{l}) = \int_{[-\pi, \pi]^d} \frac{d^d \vec{k}}{(2\pi)^d} e^{-i\vec{k} \cdot \vec{l}} \lambda^m(\vec{k})$

↳ this stems from the geometry of the lattice on which the r.w. takes place. In all generality, the Brillouin Zone appears here. (BZ)

c) $P(\vec{l}, \xi) \equiv \sum_n \xi^n P_n(\vec{l}) = \int_{BZ} \frac{d^d \vec{k}}{(2\pi)^d} e^{-i\vec{k} \cdot \vec{l}} \sum_{n=0}^{\infty} (\lambda(\vec{k}) \xi)^n$

Note that $|\lambda| \leq \sum_{\vec{l}} p(\vec{l}) = 1 ; \lambda(\vec{0}) = 1$

$P(\vec{l}, \xi) = \int_{BZ} \frac{d^d \vec{k}}{(2\pi)^d} e^{-i\vec{k} \cdot \vec{l}} \frac{1}{1 - \xi \lambda(\vec{k})}$

for $|\xi \lambda(\vec{k})| < 1$
i.e. $|\xi| < 1$

2) with the 1D random walk:

$\lambda(k) = p e^{ik} + q e^{-ik}$ (lattice spacing is unity)

$P(l, \xi) = \frac{1}{2\pi} \int_{-\pi}^{\pi} dk \frac{e^{-ikl}}{1 - \xi (p e^{ik} + q e^{-ik})}$

Case $l > 0$; change $k \rightarrow -k$ and then use $z = e^{ik}$; z runs over the unit circle \mathcal{C}

$P(l, \xi) = \frac{1}{2\pi} \oint_{\mathcal{C}} \left(\frac{dz}{iz} \right) \frac{z^l}{1 - \xi (p z^{-1} + q z)}$

$= -\frac{1}{2i\pi} \oint_{\mathcal{C}} \frac{z^l}{z^2 - \frac{1}{\xi q} z + \frac{p}{q}}$

and we will use the residue theorem

Roots of $A(z) = \xi q z^2 - z + \xi p ; \Delta = 1 - 4\xi^2 p q > 0$

$z_{\pm} = \frac{1 \pm \sqrt{1 - 4\xi^2 p q}}{2\xi q}$

Discrete random variables (5).

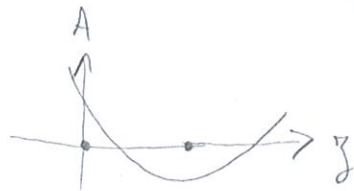
$$z_+ z_- = \frac{p}{q} > 0$$

the 2 roots have the same sign

$$A(0) = \sum p > 0$$

$$A(1) = \sum (p+q) - 1 \leq 0$$

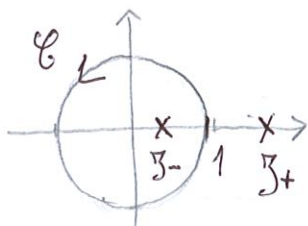
$$\Rightarrow 0 < z_- < 1 < z_+$$



of turns made by \mathcal{C} around z_k

and we apply the residue theorem:

$$\oint_{\mathcal{C}} f(z) dz = 2i\pi \sum_{z_k, \text{ residues}} \text{Res}(f, z_k) \overbrace{\text{Ind}_{\mathcal{C}}(z_k)}^{\text{around } z_k}$$



$$P(l, z) = - \frac{z^l}{zq(z_- - z_+)}$$

Only the residue at z_- does matter. Note that we took $l \geq 0$ hence z^l is not singular for $z=0$

$$P(l, z) = + \left(\frac{1 - \sqrt{1 - 4pqz^2}}{2zq} \right)^l \frac{z^l}{zq \cdot 2\sqrt{1 - 4pqz^2}} ; l \geq 0$$

For getting the result for $l < 0$, we need to exchange p and q . Thus

$$P(l, z) = \begin{cases} \left(\frac{1}{2zq} \right)^l \frac{(1 - \sqrt{1 - 4pqz^2})^l}{\sqrt{1 - 4pqz^2}} & l \geq 0 \\ \left(\frac{1}{2zp} \right)^{|l|} \frac{(1 - \sqrt{1 - 4pqz^2})^{|l|}}{\sqrt{1 - 4pqz^2}} & l \leq 0 \end{cases}$$

For $l=0$: $= \frac{1}{\sqrt{1 - 4pqz^2}}$, OK, see (B-5) $\Rightarrow R(0) = 1 - \frac{1}{\sqrt{1 - 4pqz^2}}$
 $P(0,1) = 1 - |p-q|$

Then, for $l \neq 0$: $P(l; z) = F(l; z) P(0, z)$, see (B-3)

$$F(l; z) = \frac{P(l; z)}{P(0; z)} = \begin{cases} \left(\frac{1}{2zq} \right)^l (1 - \sqrt{1 - 4pqz^2})^l & l \neq 0 \\ \left(\frac{1}{2zp} \right)^{|l|} (1 - \sqrt{1 - 4pqz^2})^{|l|} & l \neq 0 \end{cases}$$

(OK)

and we remember that $R(l) = R(D_0 + l | D_0) \equiv \sum_n F_n(D_0 + l | D_0) = F(l; \xi = 1)$

$$\begin{cases} R(l) \stackrel{l > 0}{=} \left(\frac{1}{2q}\right)^l (1 - |p - q|)^l \\ R(l) \stackrel{l < 0}{=} \left(\frac{1}{2p}\right)^{|l|} (1 - |p - q|)^{|l|} \end{cases}$$

Summary: take $p > q$

$$l > 0 \rightarrow R(l) = \left(\frac{1}{2q}\right)^l (1 - p + q)^l = 1$$

$$l = 0 \rightarrow R(0) = 1 - |p - q|$$

$$l < 0 \rightarrow R(l) = \left(\frac{1}{2p}\right)^{|l|} (1 + q + p)^l = \left(\frac{q}{p}\right)^{|l|}$$

Note that $R(1) = R(\rightarrow)$; $R(-1) = R(\leftarrow)$ found above.

3^o Polya theorem

Each lattice point has $\frac{1}{2d}$ neighbours onto which the walker can jump, with equiprobability: the structure factor of the walk then reads

$$\begin{aligned} \lambda(\vec{k}) &= \frac{1}{2d} \sum_{j=1}^d \left(e^{i k_j} + e^{-i k_j} \right) \quad \text{where } k_j \text{ is the } j^{\text{th}} \text{ Cartesian} \\ &= \frac{1}{d} \sum_{j=1}^d \cos(k_j) \quad \text{coordinate of } \vec{k} \end{aligned}$$

a) Here again, we have

$$P(\vec{r}, \xi) = \int_{\text{BZ}} \frac{d\vec{k}}{(2\pi)^d} \frac{e^{-i \vec{k} \cdot \vec{r}}}{1 - \xi \sum_{j=1}^d \cos(k_j)}$$

and the Brillouin Zone (BZ) is cubic: $[-\pi, \pi]^d$

The trick here is to exponentiate:

$$\frac{1}{1 - \xi \sum_{j=1}^d \cos(k_j)} = \int_0^\infty dt e^{-t \left[1 - \xi \sum_{j=1}^d \cos(k_j) \right]}$$

Discrete random variables (6)

$$P(\vec{r}, \vec{s}) = \frac{1}{(2\pi)^d} \int_0^\infty dt \int_{\mathbb{B}^d} d\vec{k} e^{-t} e^{\frac{\xi t}{d} \sum_{j=1}^d \cos(k_j) - i \sum_{j=1}^d k_j \cdot \ell_j}$$

$$= \frac{1}{(2\pi)^d} \int_0^\infty dt e^{-t} \int_{\mathbb{B}^d} d\vec{k} \prod_{j=1}^d e^{\frac{\xi t}{d} \cos k_j - i k_j \cdot \ell_j}$$

$$= \int_0^\infty dt e^{-t} \prod_{j=1}^d \int_0^{2\pi} \frac{dk}{2\pi} \exp\left(\frac{\xi t}{d} \cos k\right) \cos(k \ell_j)$$

↳ the function integrated is 2π -periodic
 $\Rightarrow \int_0^{2\pi} = \int_{-\pi}^{\pi}$

$$= \int_0^\infty dt e^{-t} \prod_{j=1}^d \int_0^{\pi} \frac{dk}{\pi} e^{(\xi t/d) \cos k} \cos(k \ell_j) \quad \text{thus } \gamma = \frac{\xi t}{d}$$

$$I_{\gamma} \left(\frac{\xi t}{d} \right) = I_{|\ell_j|} \left(\frac{\xi t}{d} \right)$$

$$P(\vec{r}, \vec{s}) = \int_0^\infty dt e^{-t} \prod_{j=1}^d I_{|\ell_j|} \left(\frac{\xi t}{d} \right)$$

f) We have shown that the walk is recurrent (ie the return proba to the starting site is 1) iff $P(\vec{0}, \vec{s}=1) = +\infty$. We thus have to see when the above integral diverges.

$$P(\vec{0}, \vec{s}=1) = \int_0^\infty dt e^{-t} \left[I_0 \left(\frac{\xi t}{d} \right) \right]^d$$

$$e^{-t} \left[I_0 \left(\frac{\xi t}{d} \right) \right]^d \underset{t \rightarrow \infty}{\sim} e^{-t} \left[\frac{e^{\xi t/d}}{\sqrt{2\pi \xi t/d}} \right]^d \underset{t \rightarrow \infty}{\sim} \frac{1}{(2\pi \xi t/d)^{d/2}} e^{-t} e^{-d/2}$$

Hence $P(\vec{0}, 1) = \infty \Leftrightarrow \frac{d}{2} \leq 1$

$\Leftrightarrow d \leq 2$: Polya's theorem

The r.w is recurrent in dimensions 1 and 2, and not recurrent (transient) for dimension 3, 4, ...