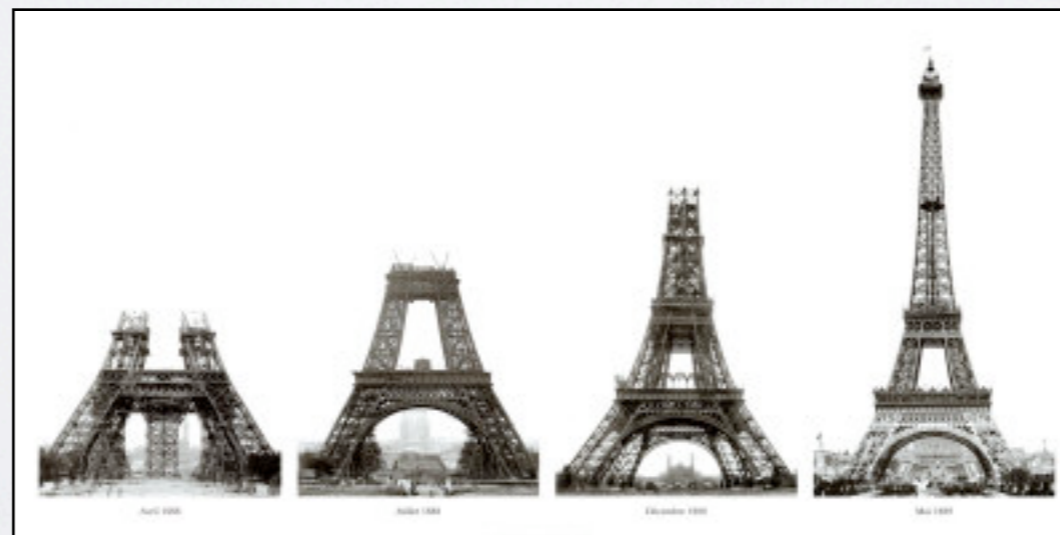


RANDOM MATRIX THEORY: OLD TRICKS FOR NEW DOGS

Pierpaolo Vivo
(LPTMS - CNRS - Paris XI)





random matrix theory biology



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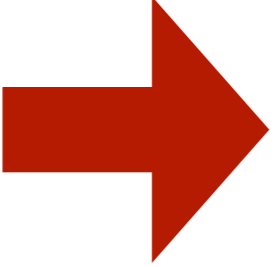
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Raffaello **Potestio**,1 Fabio Caccioli,1, 2 and Pierpaolo Vivo3. 1SISSA ... 3ICTP - Abdus Salam International Centre for Theoretical Physics, Strada Costiera 11 - 34151 Trieste, Italy ... movements of **proteins**, are investigated through the prism of **Random Matrix Theory**. ... portance for **biological** aspects ranging from structural ...

[PDF] [Article - Nano Biomedicine and Engineering](#)

nanobe.org/index.php?journal=nbe&page=article&op...

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In this work, we propose an original method, the **Random matrix theory** (RMT)-based hierarchical clustering method, to identify functional gene ... in many **biological** networks, determined the transition the decomposition of extracellular matrix **proteins**) [34]. **Potestio** R, Caccioli F, Vivo P. Random Matrix Approach ...

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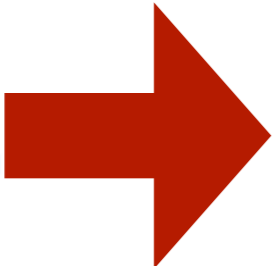
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Random Matrix approach to collective behavior and bulk universality in protein dynamics

Raffaello Potestio,¹ Fabio Caccioli,^{1,2} and Pierpaolo Vivo³

¹SISSA - Scuola Internazionale Superiore di Studi Avanzati, Via Beirut 2/4 - 34151 Trieste, Italy

²Istituto Nazionale di Fisica Nucleare, sezione di Trieste, Italy

³ICTP - Abdus Salam International Centre for Theoretical Physics, Strada Costiera 11 - 34151 Trieste, Italy

(Dated: January 12, 2010)

Covariance matrices of amino acid displacements, commonly used to characterize the large-scale movements of proteins, are investigated through the prism of Random Matrix Theory. Bulk universality is detected in the local spacing statistics of noise-dressed eigenmodes, which is well described by a Brody distribution with parameter $\beta \simeq 0.8$. This finding, supported by other consistent indicators, implies a novel quantitative criterion to single out the collective degrees of freedom of the protein from the majority of high-energy, localized vibrations.

OUTLINE

First Part: Old Tricks

- i) The old days... RMT in nuclear physics
- ii) 4 applications
 - Riemann hypothesis
 - Vicious brownian walkers
 - Covariance matrices of financial data
 - The longest increasing subsequence problem

Second Part: New Dogs

- i) Rare events and linear statistics
- ii) How many eigenvalues of a random matrix are positive?

Why are random matrix eigenvalues cool?

Message

- ❖ Ingredient: Take Any important mathematics
- ❖ Then Randomize!
- ❖ This will have many applications!

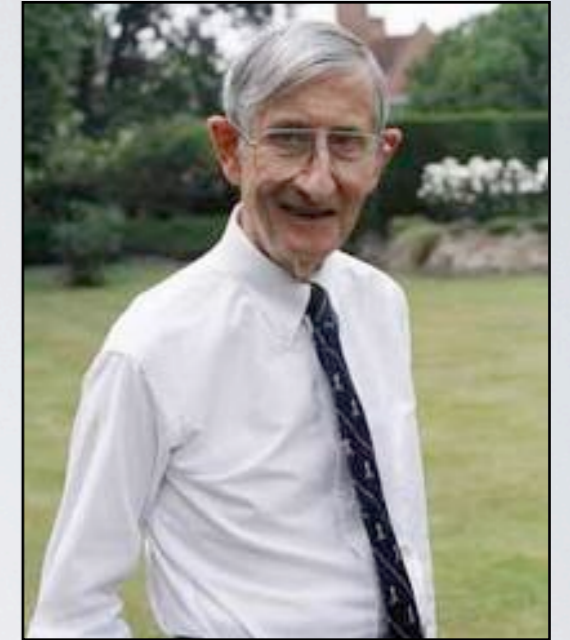
from a talk by Alan Edelman (MIT)



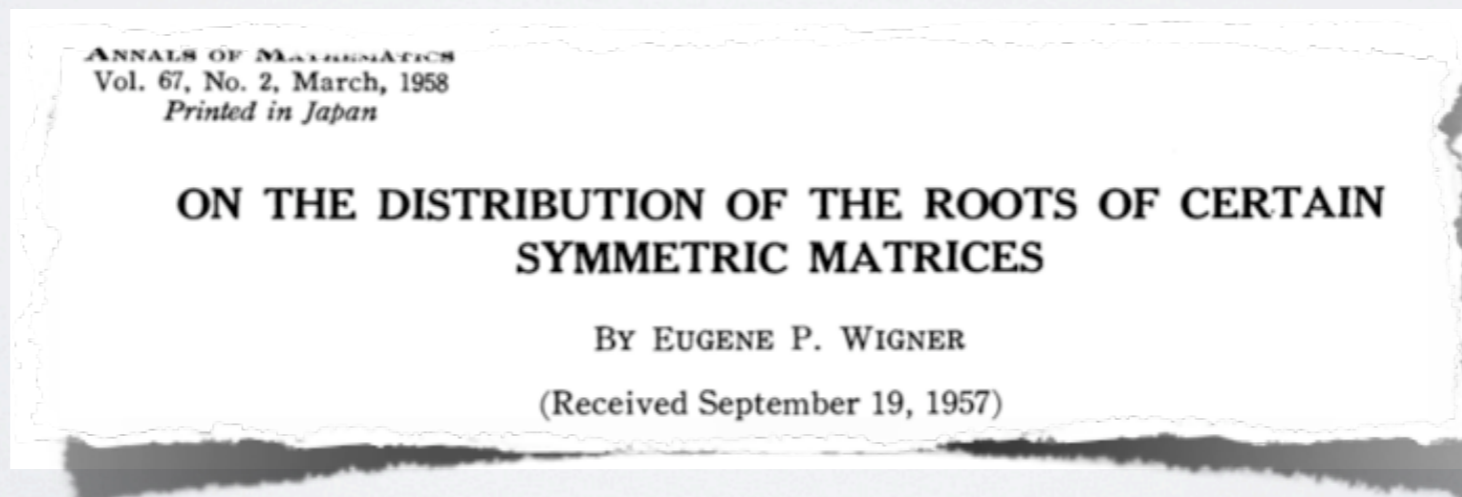
Eugene Wigner

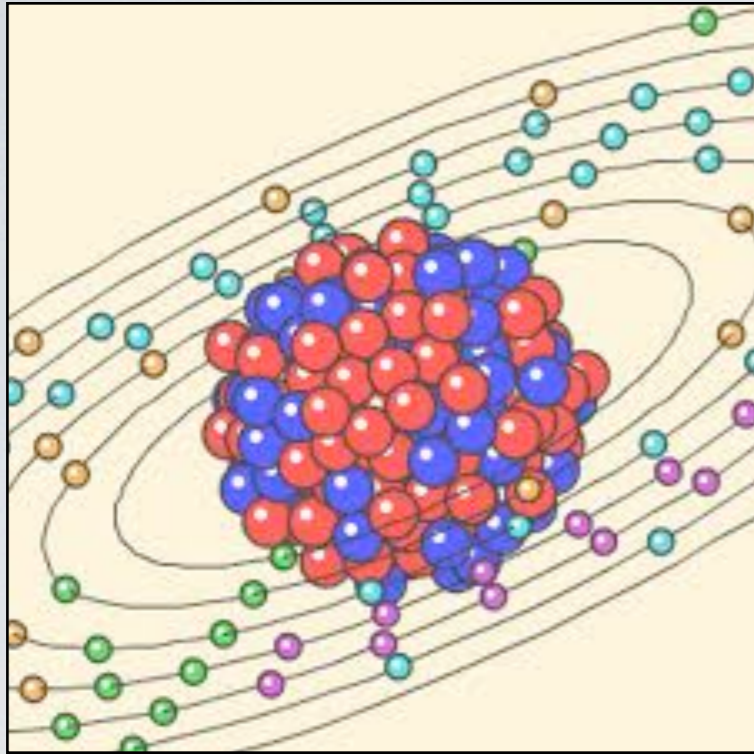


John Wishart



Freeman Dyson





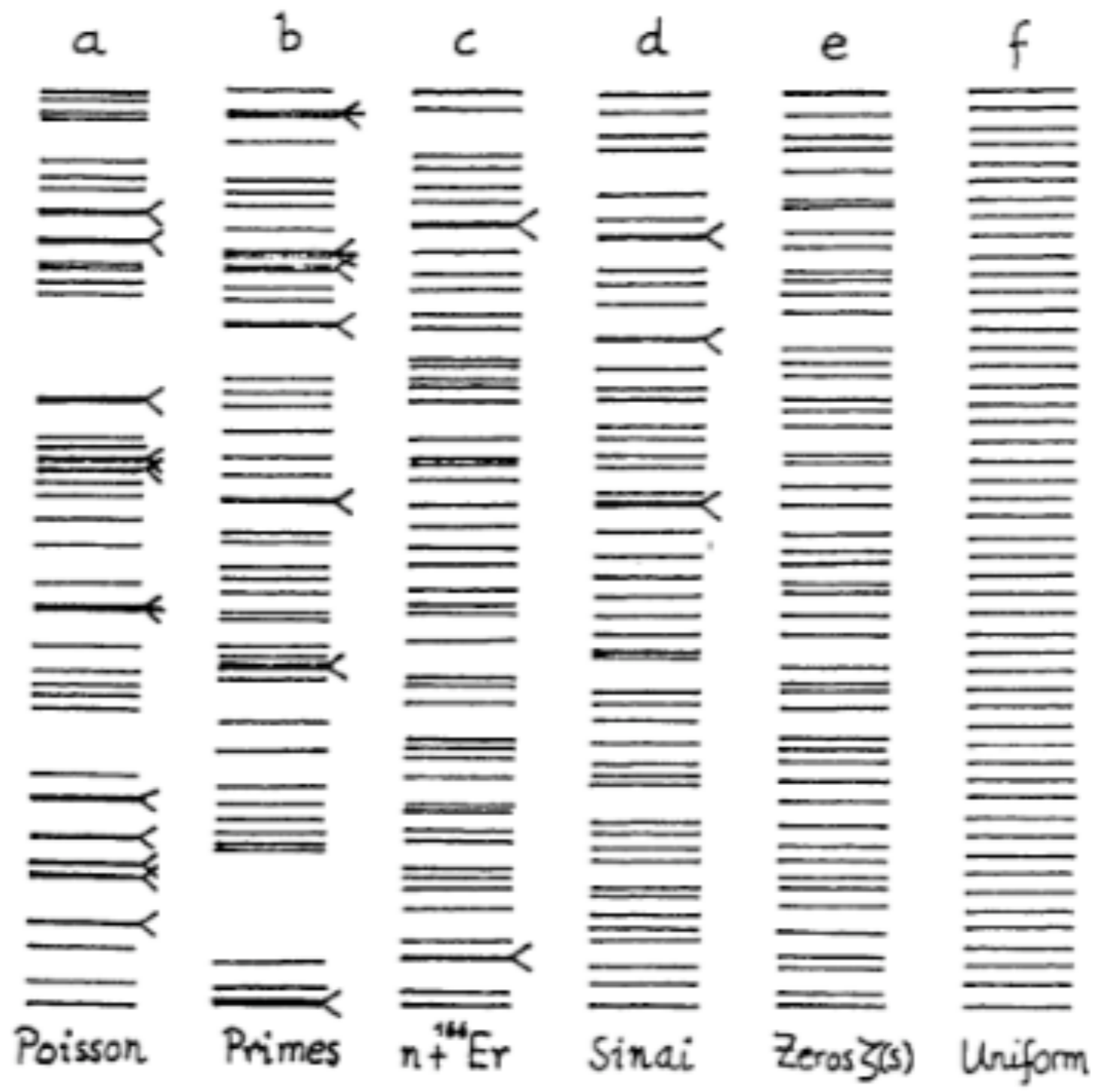
Hamiltonian (total energy) of heavy nuclei: hopeless task!

BUT.....

The Hamiltonian in a given basis is just a **HUGE** matrix....



Idea: take the matrix entries **at random...**

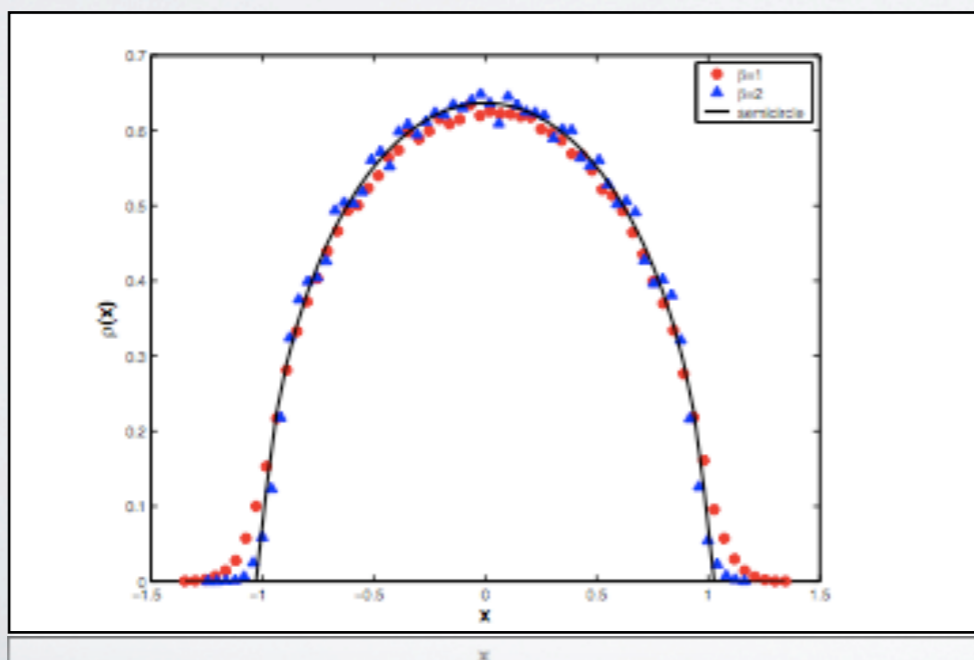


Random Matrix Theory = Randomness + Symmetry

$N = 5$

$$\begin{pmatrix} 0.5377 & 0.2631 & -1.8044 & 0.3286 & 0.4951 \\ 0.2631 & -0.4336 & 1.6888 & 1.7271 & 0.7810 \\ -1.8044 & 1.6888 & 0.7254 & 0.7133 & 0.7160 \\ 0.3286 & 1.7271 & 0.7133 & 1.4090 & 1.5237 \\ 0.4951 & 0.7810 & 0.7160 & 1.5237 & 0.4889 \end{pmatrix}$$

$$\vec{\lambda} = [-2.4341 \quad -0.8386 \quad -0.5203 \quad 2.2594 \quad 4.2610]$$

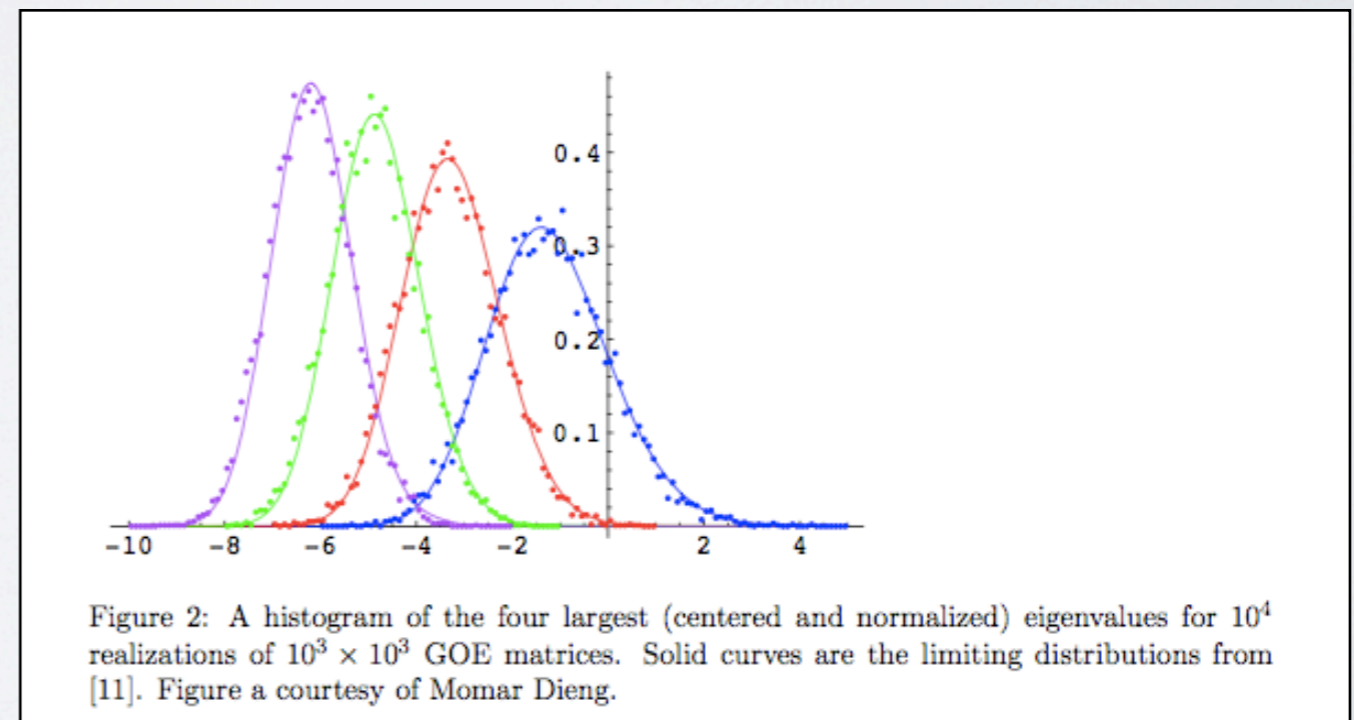
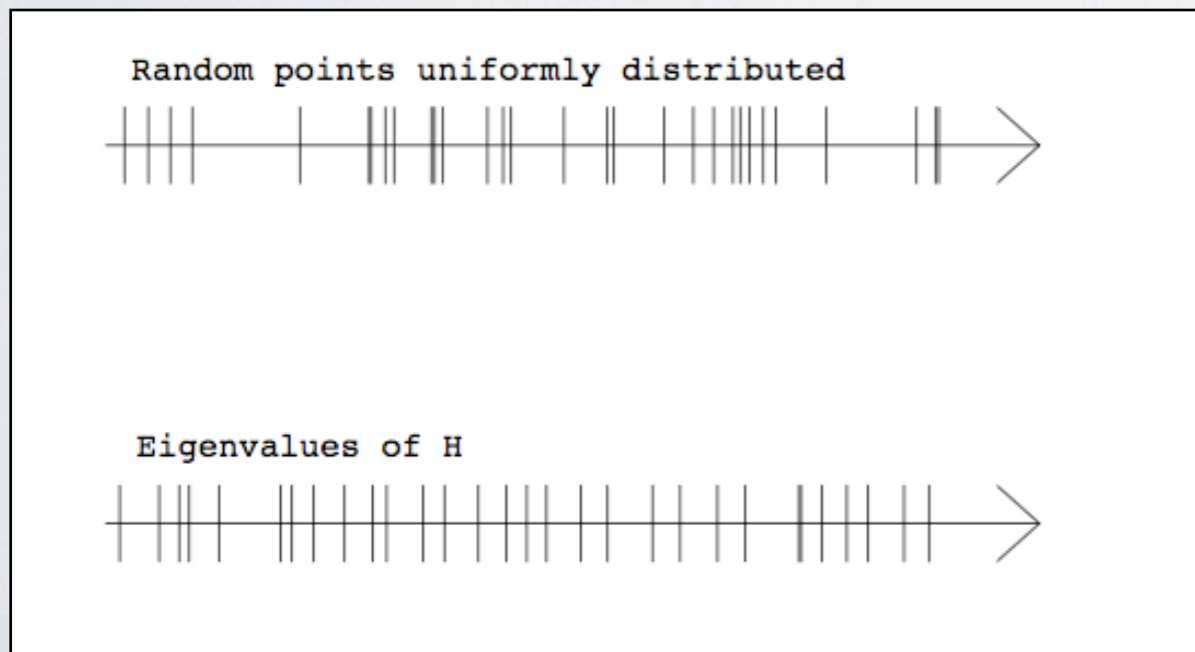


Semicircle Law

$$\rho(\lambda) = \frac{1}{2\sqrt{N}} f\left(\frac{\lambda}{2\sqrt{N}}\right)$$
$$f(x) = \frac{2}{\pi} \sqrt{1 - x^2}$$

Typical questions:

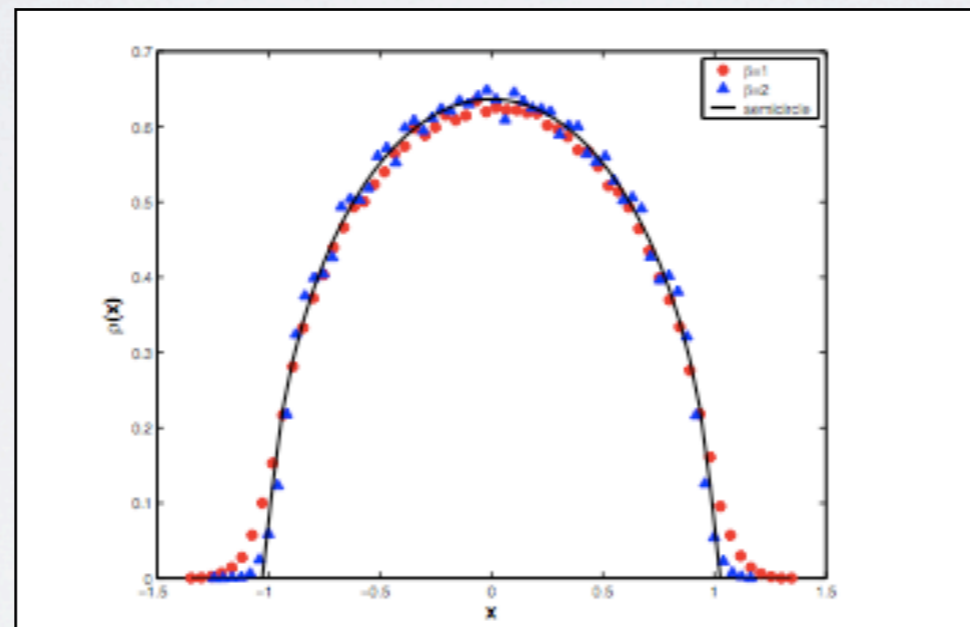
- (i) Density of eigenvalues
- (ii) Gaps between adjacent eigenvalues
- (iii) Distribution of individual eigenvalues (e.g. largest)
- (iv) Probability of **rare** events in **linear statistics**



$$\mathbb{P}_{\beta,N}(x_1, \dots, x_N) = C_{N,\beta} \prod_{1 \leq i < j \leq N} |x_i - x_j|^\beta \prod_{i=1}^N e^{-\beta x_i^2/2}, \quad \beta = 1, 2, 4,$$

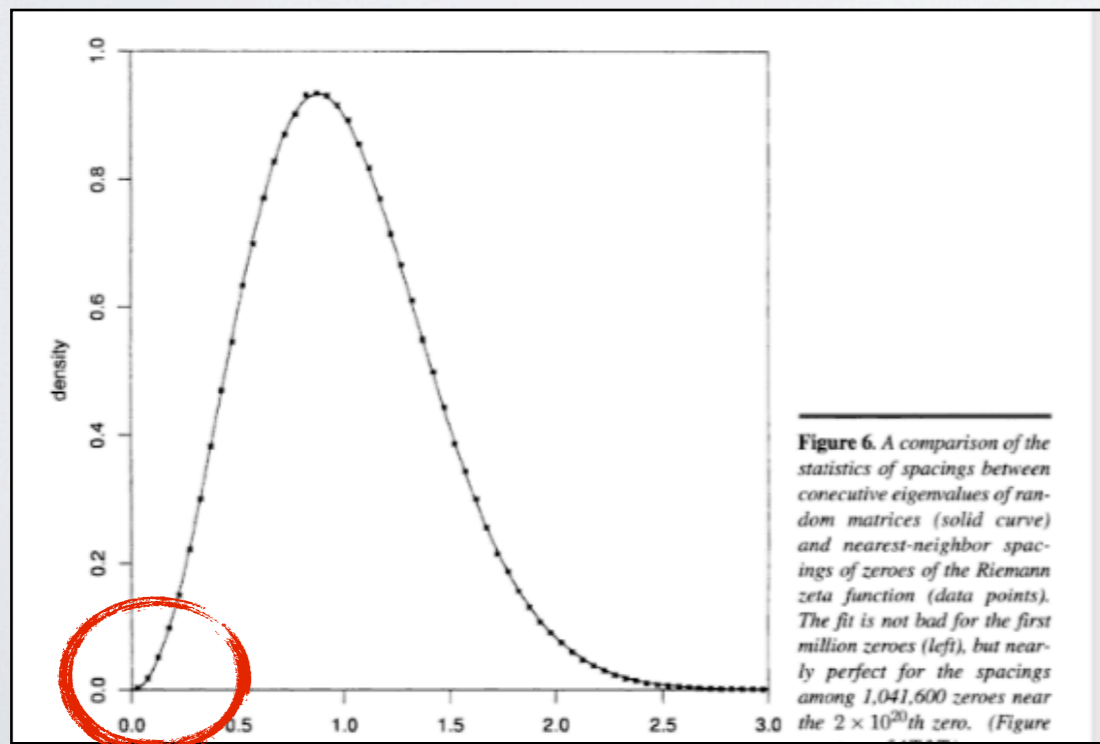
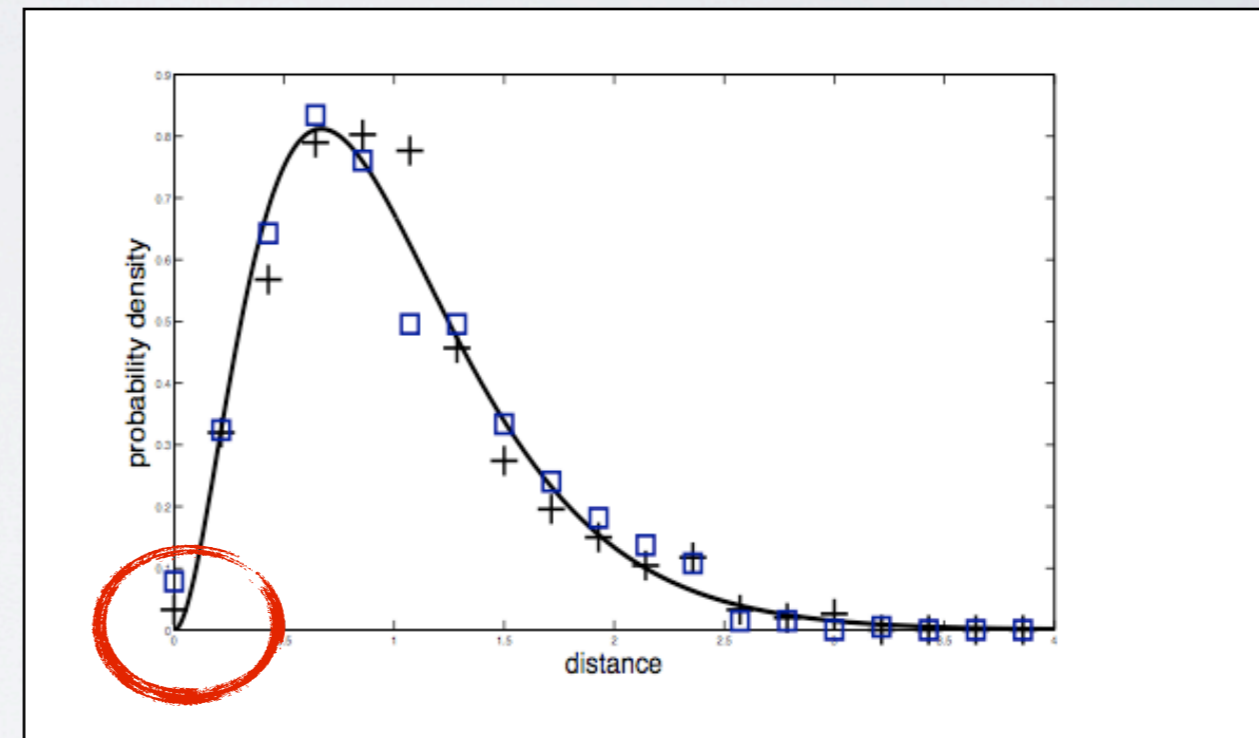
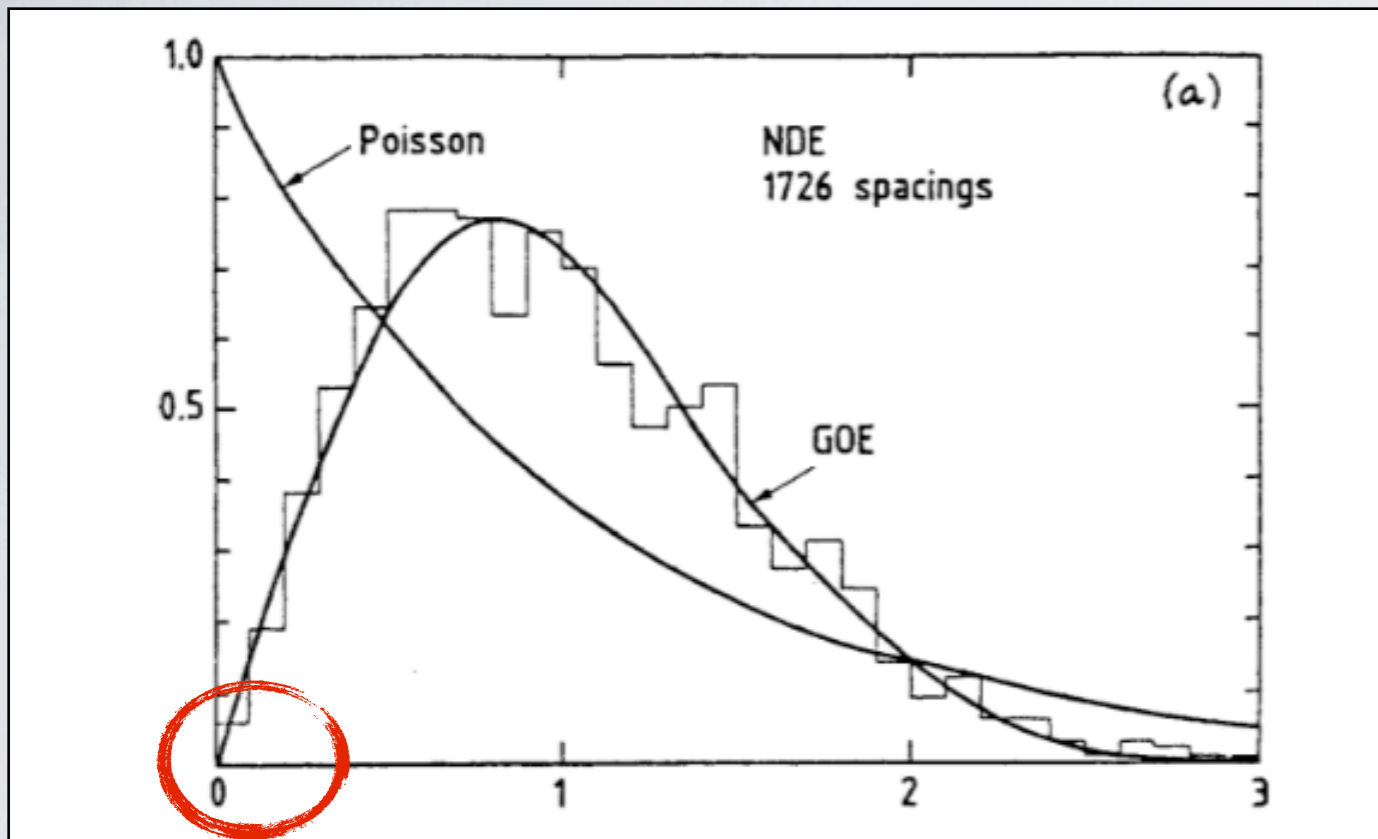
Level Repulsion

Confinement



Strongly Correlated Random Variables!!

Level Spacings: universality



$$P(s) \propto s^\beta e^{-s^2}$$

Proceedings of the 3rd Workshop on Quantum Chaos and Localisation Phenomena
Warsaw, Poland, May 25–27, 2007

Parking in the City

P. ŠEBA^{a,b,c}

^aUniversity of Hradec Králové, Hradec Králové, Czech Republic



Physica A: Statistical Mechanics and its Applications

Volume 346, Issues 3–4, 15 February 2005, Pages 621–630



Modelling the gap size distribution of parked cars

S. Rawal, G.J. Rodgers  

Department of Mathematical Sciences, Brunel University, Uxbridge, Middlesex UB8 3PH, UK

Received 19 July 2004. Available online 15 September 2004.

Modelling gap-size distribution of parked cars using random-matrix theory

A.Y. Abul-Magd

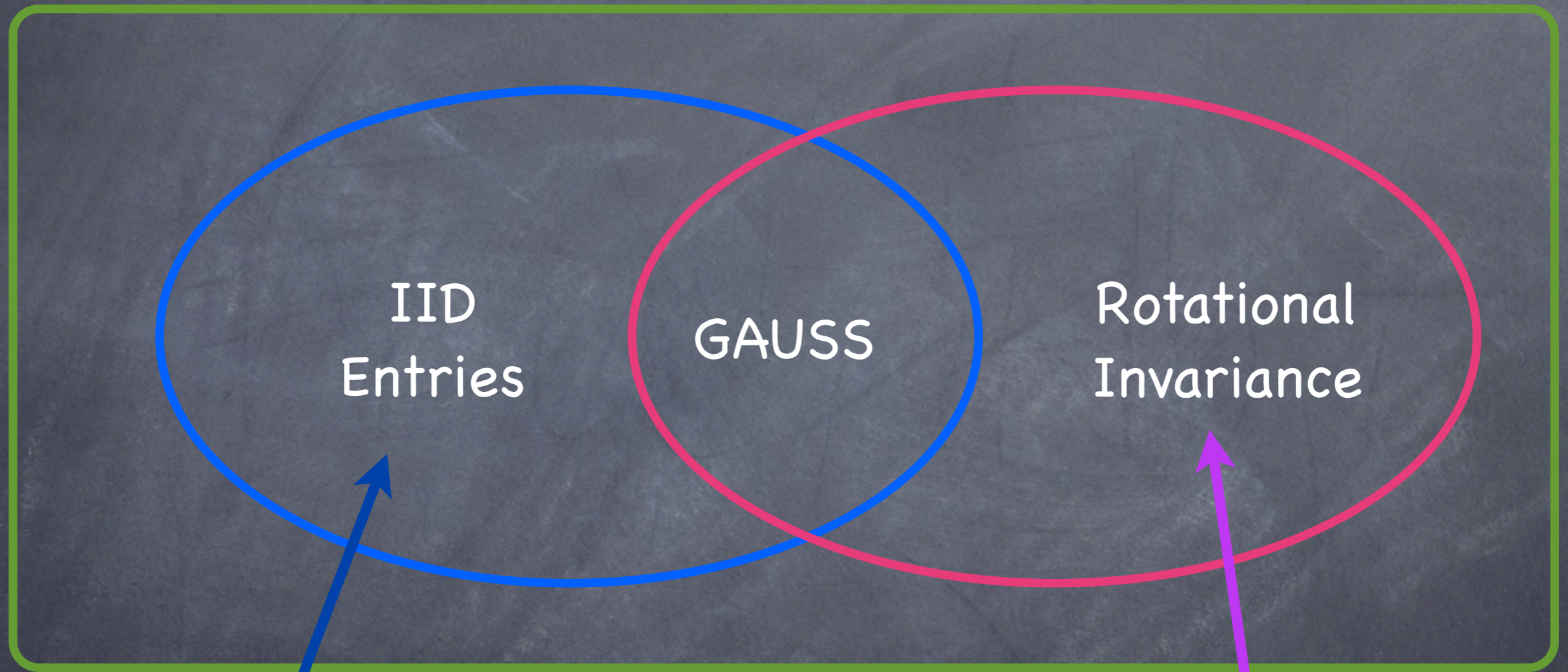
Department of Mathematics, Faculty of Science, Zagazig University, Zagazig, Egypt

We apply the random-matrix theory to the car-parking problem. For this purpose, we adopt a Coulomb gas model that associates the coordinates of the gas particles with the eigenvalues of a random matrix. The nature of interaction between the particles is consistent with the tendency of the drivers to park their cars near to each other and in the same time keep a distance sufficient for manoeuvring. We show that the recently measured gap-size distribution of parked cars in a number of roads in central London is well represented by the spacing distribution of a Gaussian unitary ensemble.

PACS: 05.40; 05.20.Gg; 02.50.r; 68.43.-h

Keywords: Car parking; Coulomb gas; Gaussian unitary ensemble

The fundamental diagram (not in textbooks)



$$\mathcal{P}(x_{11}, \dots, x_{NN}) = \prod_{ij} p(x_{ij})$$

$$\mathcal{P}(\mathbf{X}) = \mathcal{P}(\mathbf{U}\mathbf{X}\mathbf{U}^{-1})$$

$$\sum_{n=1}^{\infty} \frac{1}{n^s} = \prod_{p \text{ prime}} \frac{1}{1 - p^{-s}}$$

liegenden Wurzeln von $\zeta(t) = 0$ multiplicirt mit $2\pi i$. Man findet nun in der That etwa so viel reelle Wurzeln innerhalb dieser Grenzen, und es ist sehr wahrscheinlich, dass alle Wurzeln reell sind. Hiervon wäre allerdings ein strenger Beweis zu wünschen; ich habe indess die Aufsuchung desselben nach einigen flüchtigen vergeblichen Versuchen vorläufig bei Seite gelassen, da er für den nächsten Zweck meiner Untersuchung entbehrlich schien.

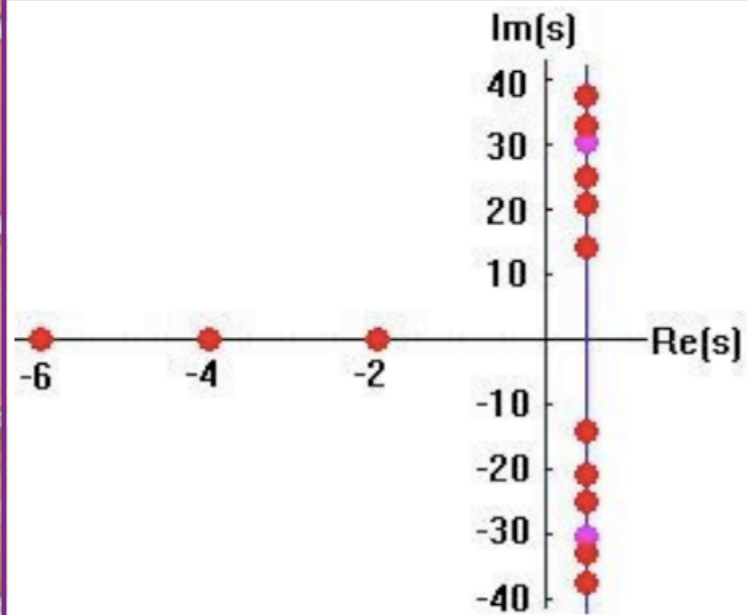


Figure 1: THE ZEROS OF THE ζ -FUNCTION.

...it is very probable that all roots are real. One would, however, wish for a strict proof of this; I have, though, after some fleeting futile attempts, provisionally put aside the search for such, as it appears unnecessary for the next objective of my investigation.

“Sometimes I think that we essentially have a complete proof of the Riemann Hypothesis except for a gap. The problem is, the gap occurs right at the beginning, and so it’s hard to fill that gap because you don’t see what’s on the other side of it.”

Montgomery's Pair Correlation Conjecture

Montgomery's pair correlation conjecture, published in 1973, asserts that the two-point correlation function $R_2(r)$ for the zeros of the Riemann zeta function $\zeta(z)$ on the critical line is

$$R_2(r) = 1 - \frac{\sin^2(\pi r)}{(\pi r)^2}.$$

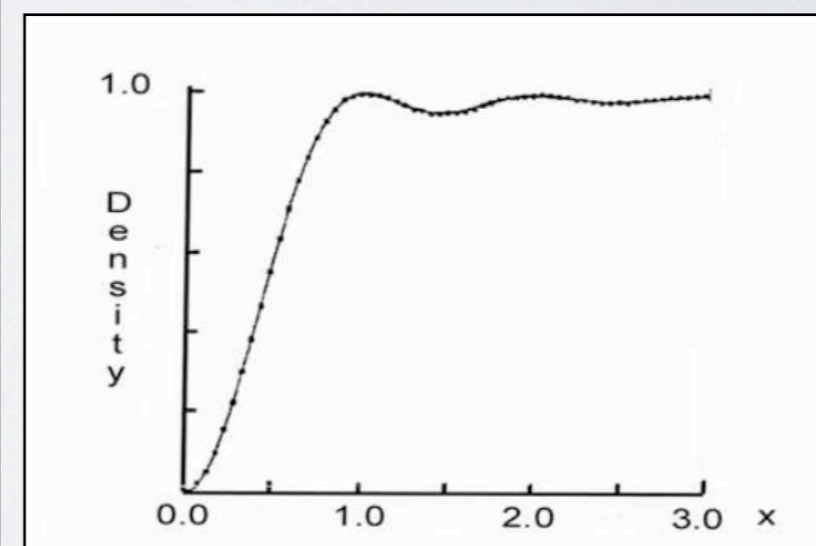
As first noted by Dyson, this is precisely the form expected for the pair correlation of random Hermitian matrices (Derbyshire 2004, pp. 287-291).

In 1972, Hugh **Montgomery**, a number theorist at the University of Michigan, was visiting the Institute for Advanced Study. **Montgomery** had been studying the distribution of zeroes of the zeta function, in hopes of gaining insight into the Riemann Hypothesis. He was able to prove that the Riemann Hypothesis had implications for the spacing of zeroes along the critical line, but his key discovery was an additional property that the zeroes seemed to have, one which implied a particularly nice formula for the average spacing between zeroes.

During tea one day at the Institute, **Montgomery** was introduced to **Dyson** and described his conjecture. **Dyson** immediately recognized it as the same result as had been obtained for random matrices.

"It just so happened that he was one of the two or three physicists in the world who had worked all of these things out, so I was actually talking to the greatest expert in exactly this!" **Montgomery** recalls.

Odlyzko's computations agree amazingly well with Montgomery's conjecture.



Simple Proof of Riemann's Hypothesis of the Zeta function

by [REDACTED], 6/20/2006

The Zeta function is defined as :

$$\zeta(s) = \sum_{k=1}^{\infty} \frac{1}{k^s} = 1 + \frac{1}{2^s} + \frac{1}{3^s} + \frac{1}{4^s} + \frac{1}{5^s} + \dots \quad s \neq 1 \quad (k=1 \rightarrow \infty)$$

Hardy, 1999 showed that :

$0 = \xi(1-s) = \xi(s)$ where $\xi(s)$ extends $\zeta(s)$ into
for all $0 < s < 1$ at the "non-trivial" zeros

$$0 = \sum_{k=1}^{\infty} \frac{1}{k^{(1-s)}} = \sum_{k=1}^{\infty} \frac{1}{k^s} \text{ for all } 0 < s < 1, \text{ positive}$$

$$0 = \sum_{k=1}^{\infty} \frac{1}{k^s} - \sum_{k=1}^{\infty} \frac{1}{k^{(1-s)}}$$

$$0 = \sum_{k=1}^{\infty} \left(\frac{1}{k^s} - \frac{1}{k^{(1-s)}} \right)$$

$$0 = \sum_{k=1}^{\infty} \left(\frac{k^{(1-s)} - k^s}{k^s k^{(1-s)}} \right)$$

$$0 = \sum_{k=1}^{\infty} \left(\frac{k^{(1-s)} - k^s}{k} \right) \Rightarrow 0 = \sum_{k=1}^{\infty} \left(\frac{0}{k} \right) \text{ if } \exists s$$

$$\Rightarrow k^{(1-s)} = k^s \Rightarrow (1-s) = s \text{ for } k=1 \rightarrow \infty, 0 < s < 1$$

substituting complex $s = (\sigma + ix)$

$$\text{for real } \sigma = 1 - \sigma \Rightarrow 2\sigma = 1$$

$$\Rightarrow \sigma = \frac{1}{2} \Rightarrow \sigma + ix = \frac{1}{2} - ix \Rightarrow \Re[s] = \frac{1}{2} \text{ for } k=1 \rightarrow \infty, 0 < s < 1, \text{ "critical strip"}$$

QED

Hardy, G. H. Ramanujan: Twelve Lectures on Subjects Suggested by His Life and Work, 3rd ed. New York: Chelsea, 1999.

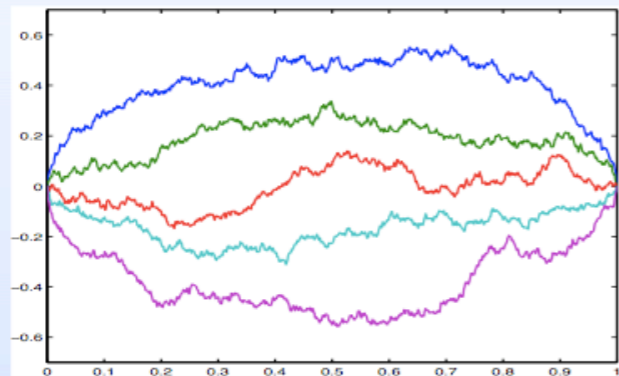
Copyright © 2006 [REDACTED]

The screenshot shows a Firefox browser window with the address bar displaying `http://arxiv.org/abs/0807.0090`. The page content includes the Cornell University Library logo, the title "A proof of the Riemann hypothesis" by Xian-Jin Li, and a notice that the paper has been withdrawn. The submission history shows four versions from July 2008. The browser's taskbar at the bottom displays various application icons and the system clock showing Friday, 11:31 PM.

Non-intersecting Brownian motion paths

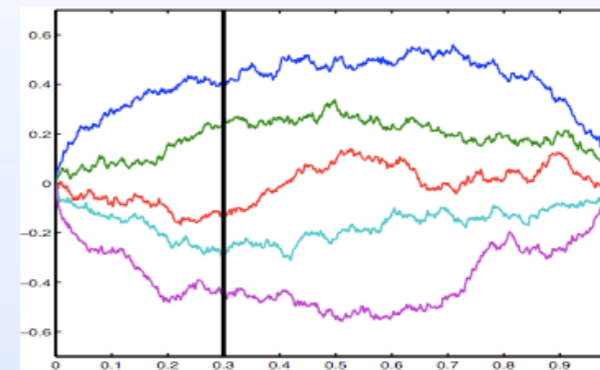
- ▲ Take n independent 1-dimensional Brownian motions with time in $[0, 1]$ conditioned so that:

- ▲ All paths start and end at the same point.
- ▲ The paths **do not intersect** at any intermediate time.



Five non-intersecting Brownian bridges

- ▲ **Remarkable fact:** At any intermediate time the positions of the paths have **exactly the same distribution** as the eigenvalues of an $n \times n$ GUE matrix (up to a scaling factor).



Positions of five non-intersecting Brownian paths behave the same as the eigenvalues of a 5×5 GUE matrix

- ▲ This interpretation is basic for the connection of random matrix theory with growth models of statistical physics.

Introduction. Since the pioneering work of de Gennes [1], followed up by Fisher [2], the subject of vicious (non-intersecting) random walkers has attracted a lot of interest among physicists. It has been studied in the context of wetting and melting [2], networks of polymers [3] and fibrous structures [1], persistence properties in nonequilibrium systems [4] and stochastic growth models [5, 6]. There also exist connections between the

PHYSICAL REVIEW E

VOLUME 52, NUMBER 6

DECEMBER 1995

Vicious walkers and directed polymer networks in general dimensions

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Department of Mathematics, Royal Holloway, University of London, Egham Hill, Egham, Surrey TW200EX, United Kingdom

A. J. Guttmann

Department of Mathematics, The University of Melbourne, Parkville, Victoria 3052, Australia

(Received 2 May 1995)

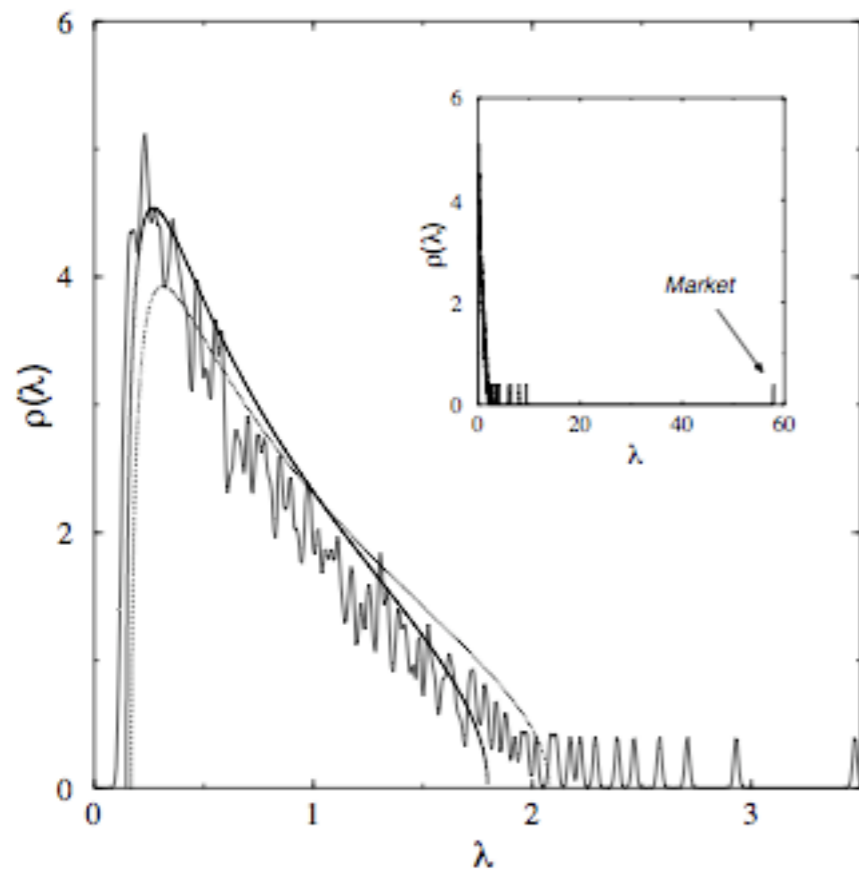
Random Covariance Matrices

$$\mathbf{X} = \begin{array}{c} \text{1} \\ \text{2} \\ \text{3} \end{array} \begin{array}{c} \text{phys.} \\ \text{math} \end{array} \begin{array}{|c|c|} \hline \mathbf{X}_{11} & \mathbf{X}_{12} \\ \hline \mathbf{X}_{21} & \mathbf{X}_{22} \\ \hline \mathbf{X}_{31} & \mathbf{X}_{33} \\ \hline \end{array} \quad \begin{array}{l} \text{in general} \\ \text{(MxN)} \end{array}$$

$$\mathbf{X}^t = \begin{array}{|c|c|c|} \hline \mathbf{X}_{11} & \mathbf{X}_{21} & \mathbf{X}_{31} \\ \hline \mathbf{X}_{12} & \mathbf{X}_{22} & \mathbf{X}_{33} \\ \hline \end{array} \quad \begin{array}{l} \text{in general} \\ \text{(NxM)} \end{array}$$

$$\mathbf{W} = \mathbf{X}^t \mathbf{X} = \begin{array}{|c|c|} \hline \mathbf{X}_{11}^2 + \mathbf{X}_{21}^2 + \mathbf{X}_{31}^2 & \mathbf{X}_{11}\mathbf{X}_{12} + \mathbf{X}_{21}\mathbf{X}_{22} + \mathbf{X}_{31}\mathbf{X}_{33} \\ \hline \mathbf{X}_{12}\mathbf{X}_{11} + \mathbf{X}_{22}\mathbf{X}_{21} + \mathbf{X}_{33}\mathbf{X}_{31} & \mathbf{X}_{12}^2 + \mathbf{X}_{22}^2 + \mathbf{X}_{33}^2 \\ \hline \end{array}$$

\rightarrow (NxN) COVARIANCE MATRIX (unnormalized)



Noise Dressing of Financial Correlation Matrices

Laurent Laloux,^{1,*} Pierre Cizeau,¹ Jean-Philippe Bouchaud,^{1,2} and Marc Potters¹

¹Science & Finance, 109-111 rue Victor Hugo, 92532 Levallois Cedex, France

²Service de Physique de l'État Condensé, Centre d'études de Saclay, Orme des Merisiers, 91191 Gif-sur-Yvette Cedex, France

(Received 15 December 1998)

time. From this point of view, it is interesting to compare the properties of an empirical correlation matrix \mathbf{C} to a null hypothesis purely *random* matrix as one could obtain from a finite time series of strictly independent assets. Deviations from the random matrix case might then suggest the presence of true information. The theory of random matri-

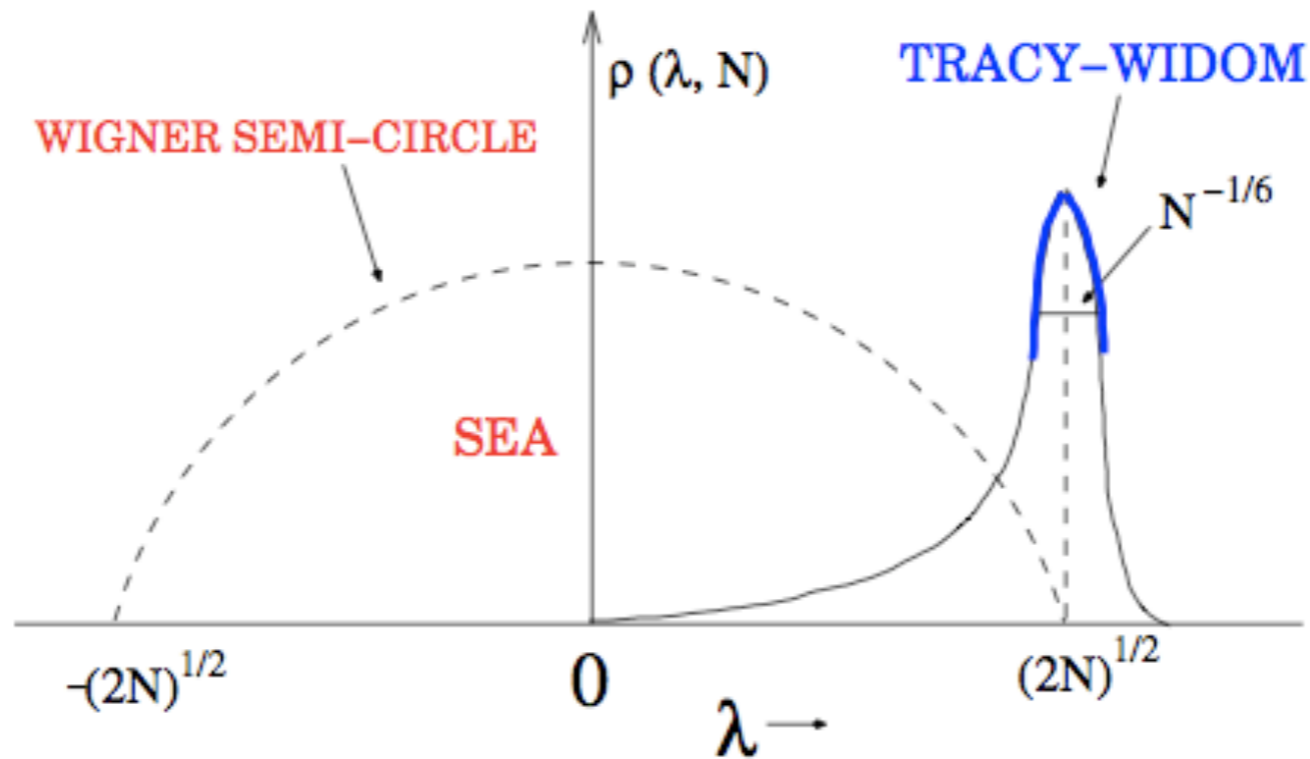
Debate: is the bulk of the stock market correlation matrix just pure noise?

A new method to estimate the noise in financial correlation matrices

Thomas Guhr¹ and Bernd Kälber^{2,3}

LARGEST EIGENVALUE

Tracy-Widom distribution for λ_{\max}

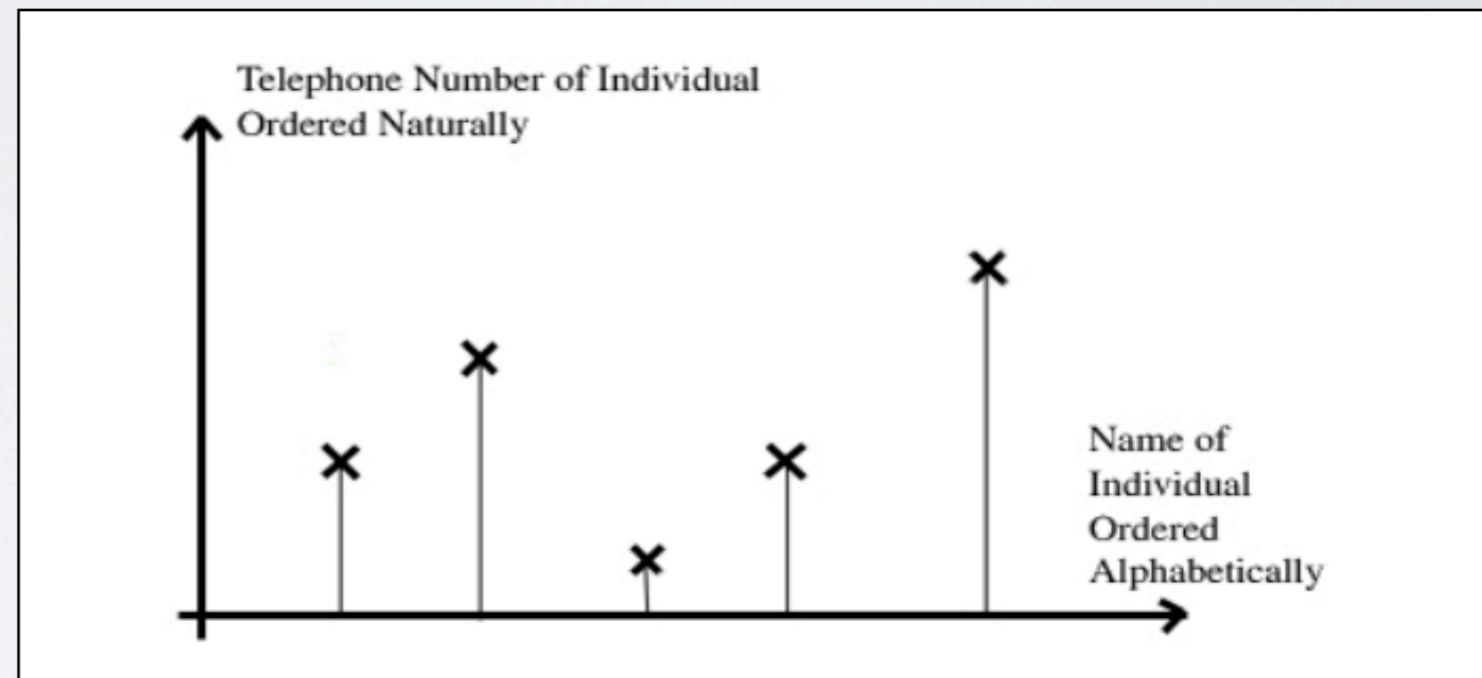


- $\langle \lambda_{\max} \rangle = \sqrt{2N}$; typical fluctuation: $|\lambda_{\max} - \sqrt{2N}| \sim N^{-1/6}$ (small)
- typical fluctuations are distributed via Tracy-Widom (1994):
- cumulative distribution:
$$\text{Prob}[\lambda_{\max} \leq t, N] \rightarrow F_{\beta} \left(\sqrt{2} N^{1/6} (t - \sqrt{2N}) \right)$$
- Prob. density (pdf): $f_{\beta}(z) = dF_{\beta}(z)/dz$
- $F_{\beta}(z) \rightarrow$ obtained from solution of Painlevé-II equation

“Are Tracy and Widom in Your Local Telephone Directory?”

Ryan Witko

Advisor: Percy Deift



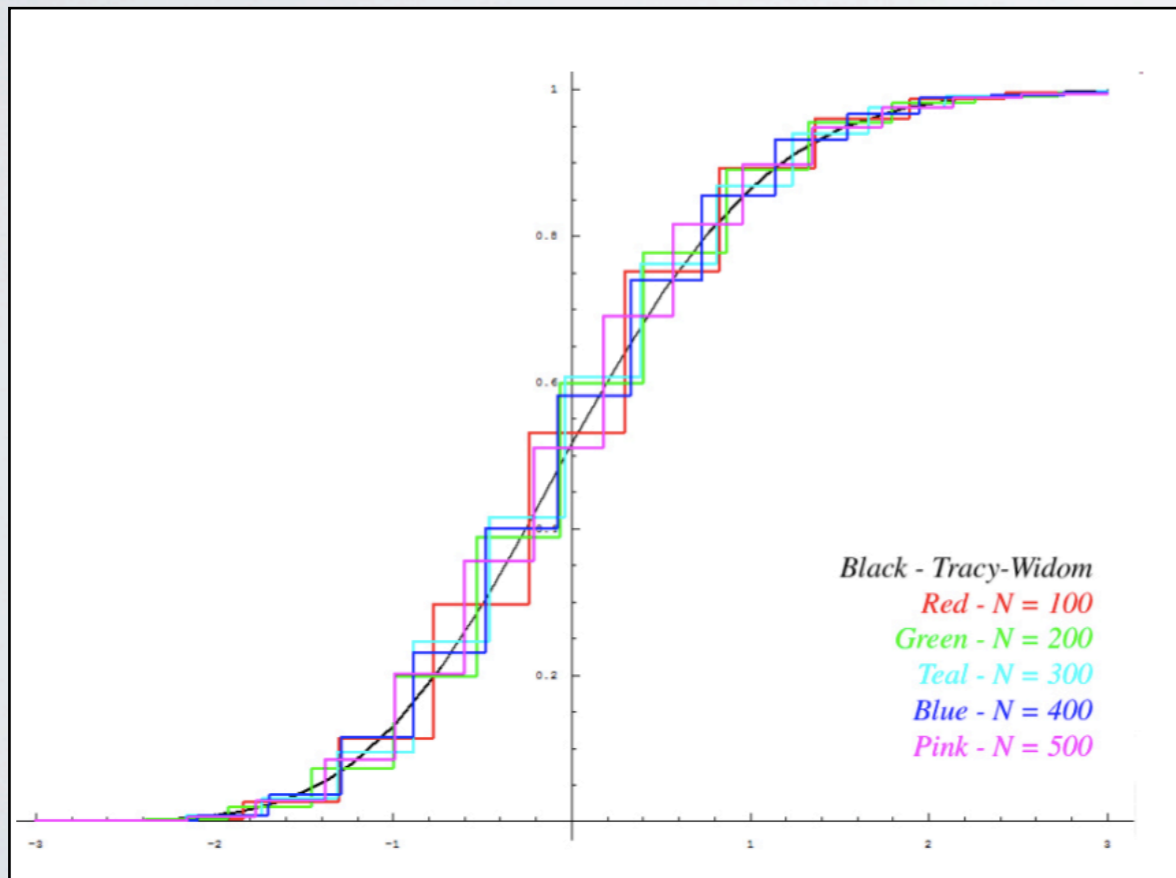
Definition:

The **longest increasing (contiguous) subsequence** of a given sequence is the subsequence of increasing terms containing the largest number of elements. For example, the longest increasing subsequence of the permutation {6, 3, 4, 8, 10, 5, 7, 1, 9, 2} is {3, 4, 8, 10}.

It can be coded in *Mathematica* as follows.

```
<<Combinatorica`  
LongestContiguousIncreasingSubsequence[p_] :=  
  Last [  
    Split[Sort[Runs[p]], Length[#1] >= Length[#2] &]  
  ]
```

[More information »](#)



We broke the 647,028 entries into successive samples each containing N entries.

647,028

Jinho Baik, Kurt Johansson and Percy Deift showed that as $N \rightarrow \infty$

(3)
$$\text{Prob} \left(\frac{\ell_N - 2\sqrt{N}}{N^{1/6}} \leq t \right) \rightarrow F(t)$$

The function $F(t)$ was shown by Craig Tracy and Harold Widom to be the distribution of the largest eigenvalue of a random matrix in the Gaussian Unitary Ensemble (GUE). It

SUMMARY

- Eigenvalues of random matrices: strongly correlated
- Level Repulsion
- Tracy-Widom distribution: analogue of Gaussian distribution for correlated random variables
- Zeros of Riemann zeta have the same statistical properties as the eigenvalues of Gaussian matrices
- Non-intersecting Brownian bridges
- Wishart matrices: covariance matrices of random data

SECOND PART

Probability of **rare** events in **linear statistics**

“Lies, damned lies, and statistics.”

A simple example of large deviation tails

- Let $M \rightarrow$ no. of heads in N tosses of an unbiased coin
- Clearly $P(M, N) = \binom{N}{M} 2^{-N}$ ($M = 0, 1, \dots, N$) \rightarrow binomial distribution

with mean = $\langle M \rangle = \frac{N}{2}$ and variance = $\sigma^2 = \langle (M - \frac{N}{2})^2 \rangle = \frac{N}{4}$

- typical fluctuations $M - \frac{N}{2} \sim O(\sqrt{N})$ are well described by the Gaussian form: $P(M, N) \sim \exp \left[-\frac{2}{N} (M - \frac{N}{2})^2 \right]$
- Atypical large fluctuations $M - \frac{N}{2} \sim O(N)$ are not described by Gaussian form
- Setting $M/N = x$ and using Stirling's formula $N! \sim N^{N+1/2} e^{-N}$ gives

$$P(M = Nx, N) \sim \exp[-N\Phi(x)] \quad \text{where}$$

$\Phi(x) = x \log(x) + (1-x) \log(1-x) + \log 2$ \rightarrow large deviation function

- $\Phi(x) \rightarrow$ symmetric with a minimum at $x = 1/2$ and for small arguments $|x - 1/2| \ll 1$, $\Phi(x) \approx 2(x - 1/2)^2$
 \rightarrow recovers the Gaussian form near the peak

LINEAR STATISTICS

 $\{x_1, \dots, x_N\}$

Random Variables

$$A = \sum_{i=1}^N f(x_i)$$

Question: what is the distribution of A for large N ?

Independent

Correlated

Central Limit Theorems

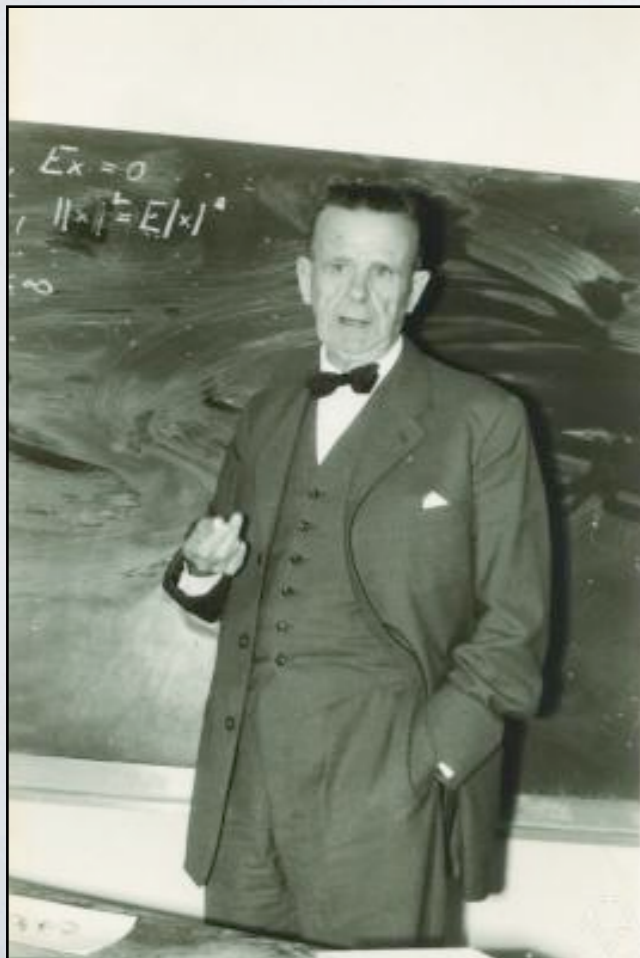
?

The first rigorous results concerning large deviations are due to the Swedish mathematician Harald Cramér, who applied them to model the insurance business. From the point of view of an insurance company, the earning is at a constant rate per month (the monthly premium) but the claims X_i come randomly.

For the company to be successful over a certain period of time (preferably many months), the total earning should exceed the total claim.

Thus to estimate the premium you have to ask the following question : "What should we choose as the premium q such that over N months the total claim $C = \sum_i X_i$ should be less than Nq ? "

Cramér gave a solution to this question for i.i.d. random variables



What if the random variables are strongly correlated?

A Nontrivial Problem

REAL SYMMETRIC MATRIX (N×N)

$$\mathbf{X} = \begin{pmatrix} \mathbf{x}_{11} & \mathbf{x}_{12} & \cdot & \cdot & \cdot & \cdot & \cdot & \mathbf{x}_{1N} \\ \mathbf{x}_{21} & \mathbf{x}_{22} & \cdot & \cdot & \cdot & \cdot & \cdot & \mathbf{x}_{2N} \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot \\ \mathbf{x}_{N1} & \cdot & \cdot & \cdot & \cdot & \cdot & \cdot & \mathbf{x}_{NN} \end{pmatrix}$$

GAUSSIAN
↓
Pr [X]
∝
 $\exp[-\frac{1}{2}\text{Tr}(\mathbf{X}^2)]$

N eigenvalues : $\lambda_1, \lambda_2, \dots, \lambda_N$
↳ strongly correlated

- $P_N = \text{Prob}[\lambda_1 \leq 0, \lambda_2 \leq 0, \dots, \lambda_N \leq 0] = \text{Prob}[\lambda_{\max} \leq 0] = ?$

[R.M. May, Nature, 238, 413 (1972)—Ecosystems]

[Cavagna et. al. 2000, Fyodorov 2004, — Glassy systems]

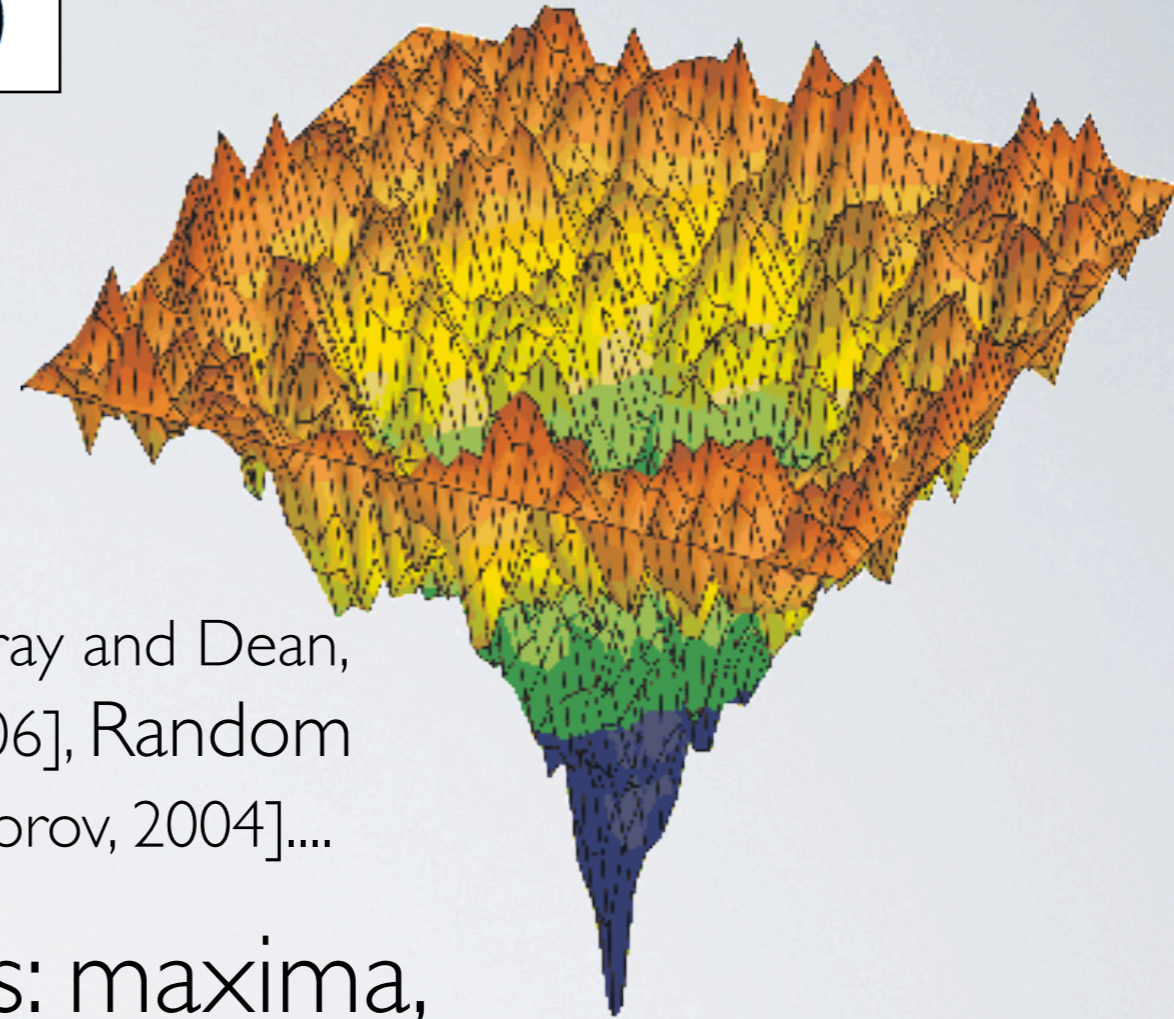
[Susskind 2003, Douglas et. al. 2004, Aazami & Easter 2006— String theory].....

A particle moving in a

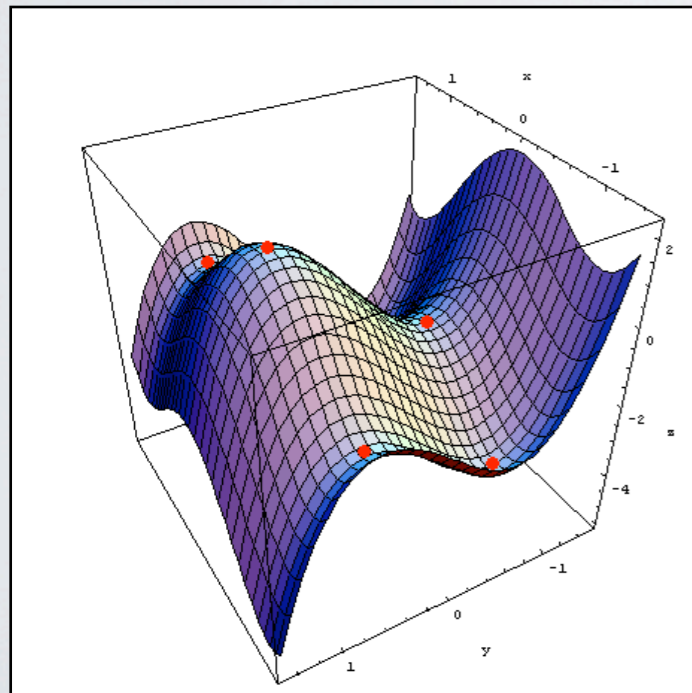
N-dim. landscape

$$V(y_1, \dots, y_N)$$

$$\frac{dy_i}{dt} = -\nabla_{y_i} V$$



- Spin and structural glasses, Gaussian fields [Bray and Dean, 2006], String landscapes [Aazami and Easter, 2006], Random Energy Landscapes and Glass Transition [Fyodorov, 2004]....



Stationary points: maxima,
minima and saddles



$$H_{i,j} = \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$$

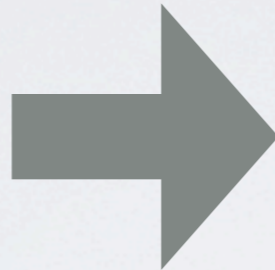
Hessian matrix

Eigenvalues of Hessian matrix determine the nature of the stationary point

RANDOM HESSIAN MODEL

Draw the elements of the Hessian matrix independently at random

$$H_{i,j} = \left[\frac{\partial^2 V}{\partial y_i \partial y_j} \right]$$



It belongs to the **GOE** of random matrices

The index distribution (number of positive eigenvalues) provides information about the **typical** stability pattern of a Random Hessian model



Most of the stationary points are saddles!

Cosmology from random multifield potentials

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Downloaded 23 January 2006

JCAP

Despite the approximation used to obtain equation (8), we have confirmed that the likelihood that all the eigenvalues of an $N \times N$ symmetric matrix have the same sign scales as e^{-cN^2} . The measured constant differs slightly from -0.25 , although given the simplicity of our approximation the agreement is perhaps surprisingly good.

- Based on numerics, Aazami & Easther (2006) predicted for large N :

$$P_N \sim \exp[-\theta N^2] \text{ with } \theta_{\text{num}} \approx 0.27$$

→ very small probability → RARE EVENT

- Exact result: $\theta = \frac{1}{4} \ln(3) = 0.274653..$ (Dean and S.M., 2006)

More generally, for $\beta = 1$ (GOE), $\beta = 2$ (GUE) and $\beta = 4$ (GSE)

$$P_N \sim \exp[-\beta\theta N^2] \text{ for large } N$$

GAUSSIAN MATRIX $N \times N$

$N = 5$

$$\begin{pmatrix} 0.5377 & 0.2631 & -1.8044 & 0.3286 & 0.4951 \\ 0.2631 & -0.4336 & 1.6888 & 1.7271 & 0.7810 \\ -1.8044 & 1.6888 & 0.7254 & 0.7133 & 0.7160 \\ 0.3286 & 1.7271 & 0.7133 & 1.4090 & 1.5237 \\ 0.4951 & 0.7810 & 0.7160 & 1.5237 & 0.4889 \end{pmatrix}$$

Real Symmetric or Complex Hermitian or Quaternion self-dual :
eigenvalues are **real**

$$\vec{\lambda} = [-2.4341 \quad -0.8386 \quad -0.5203 \quad 2.2594 \quad 4.2610]$$

\mathcal{N}_+ = number of positive eigenvalues

 The index

$$P_\beta(\lambda_1, \dots, \lambda_N)$$

Joint probability density of eigenvalues

$$\mathcal{P}(\mathcal{N}_+, N) = \int_{-\infty}^{\infty} d\lambda_1 \cdots d\lambda_N P_\beta(\lambda_1, \dots, \lambda_N) \delta\left(\mathcal{N}_+ - \sum_{i=1}^N \theta(\lambda_i)\right)$$

Probability distribution of linear statistics [Chen '94 - '98
Forrester '98
Beenakker '93
.....]

$$P_\beta(\lambda_1, \dots, \lambda_N) = \frac{1}{Z_N} e^{-\frac{\beta}{2} \sum_{i=1}^N \lambda_i^2} \prod_{j < k} |\lambda_j - \lambda_k|^\beta = \frac{1}{Z_N} e^{-\beta \mathcal{H}(\vec{\lambda})}$$

$$\mathcal{H}(\vec{\lambda}) = \frac{1}{2} \sum_{i=1}^N \lambda_i^2 - \frac{1}{2} \sum_{j \neq k} \log |\lambda_j - \lambda_k|$$

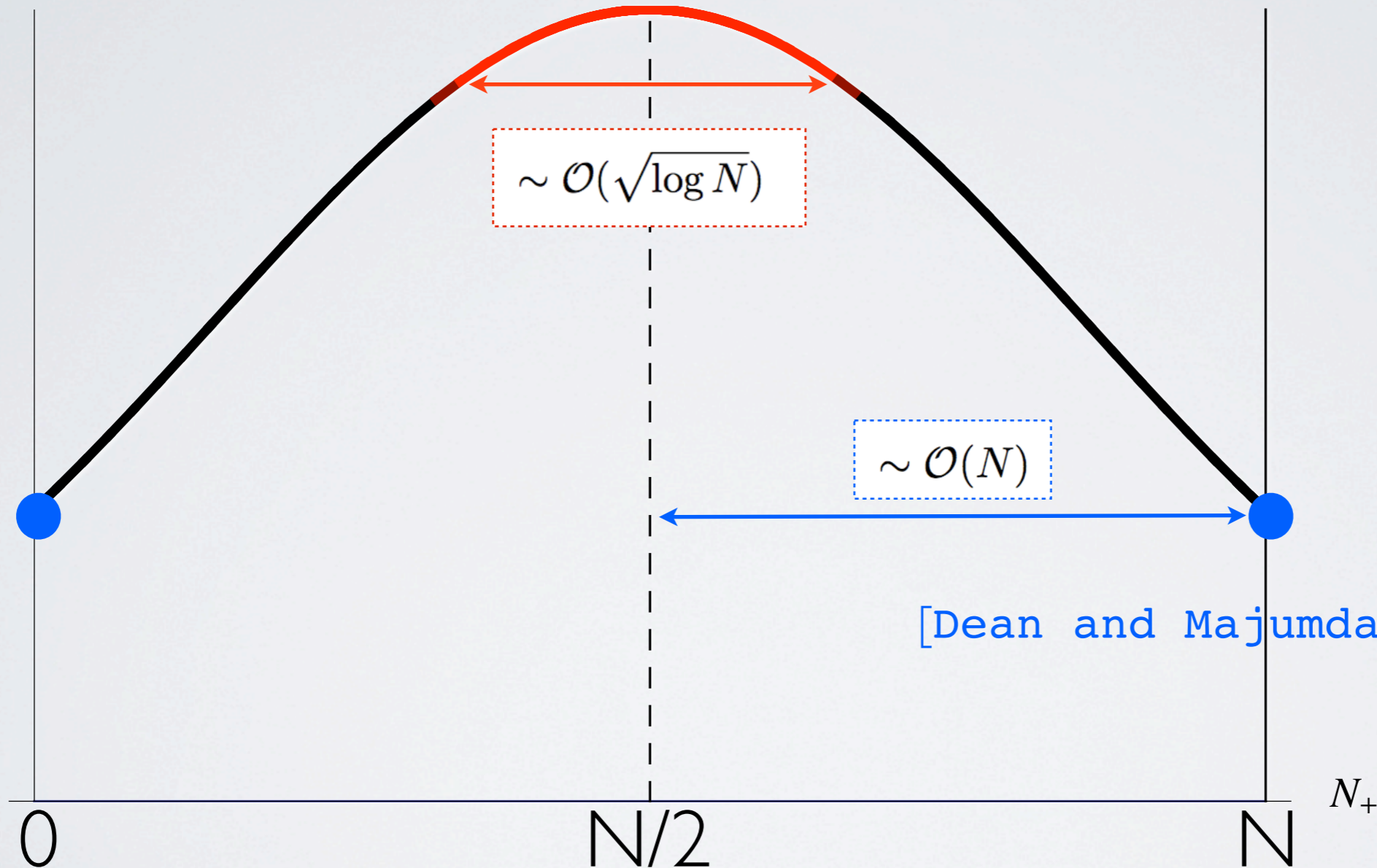
Canonical weight of an auxiliary thermodynamical system

WHAT IS KNOWN? 2 SCALES IN THIS PROBLEM

\mathcal{N}_+ = number of positive eigenvalues

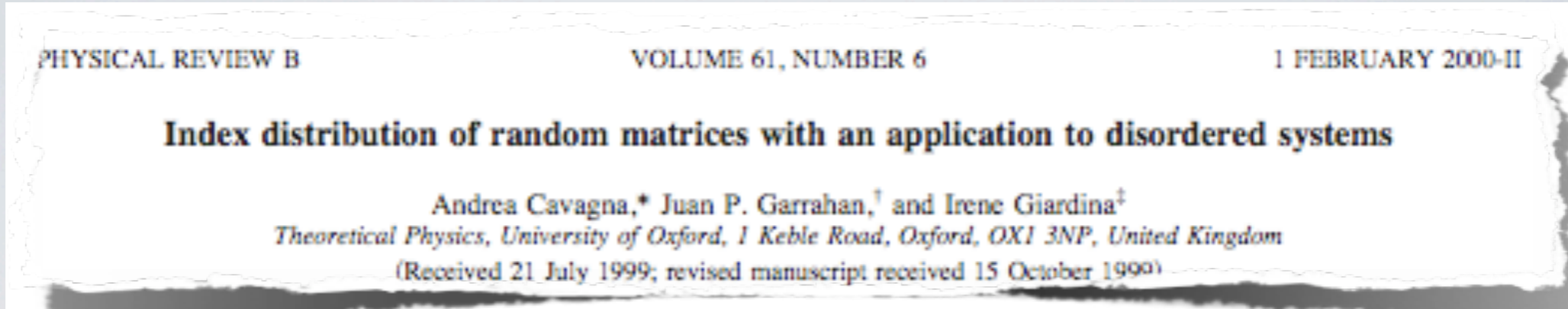
Prob[\mathcal{N}_+]

[Cavagna *et al.* (2000)]



TYPICAL VS. ATYPICAL FLUCTUATIONS: A PUZZLE?

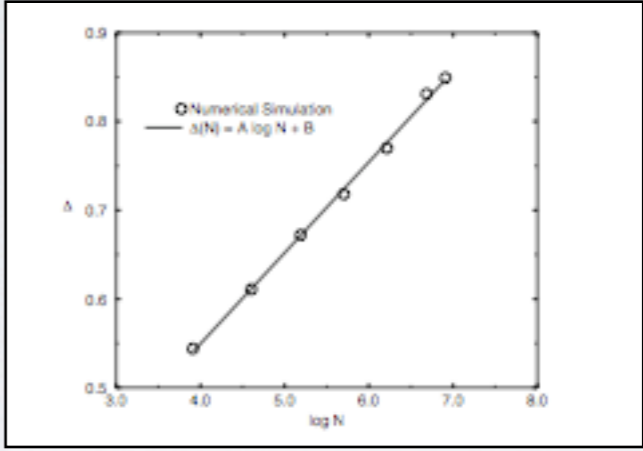
Peak



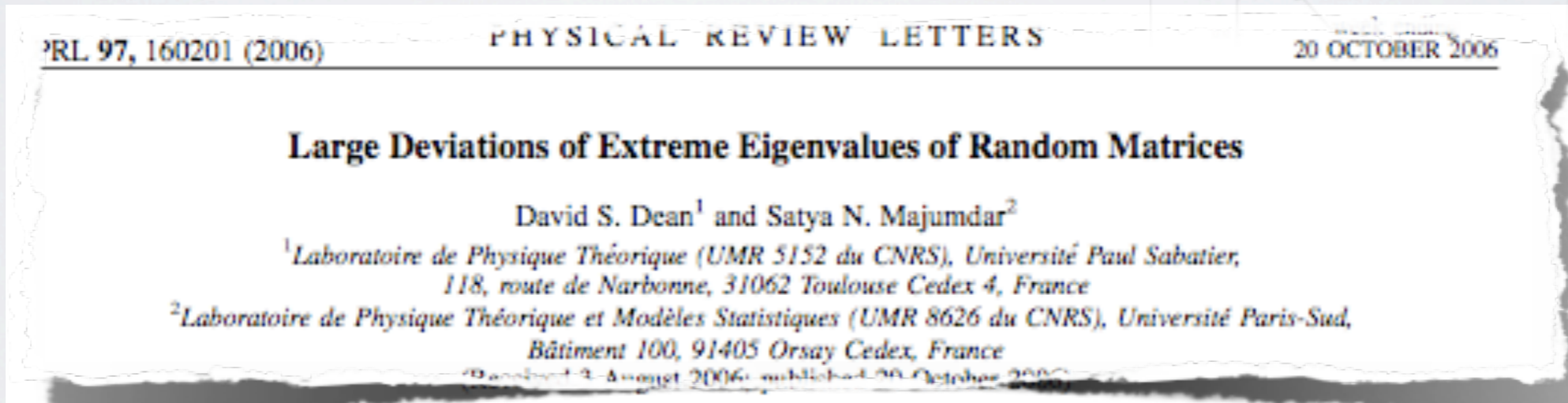
$$c = \frac{\mathcal{N}_+}{N}$$

$$\mathcal{P}(\mathcal{N}_+ = cN, N) \simeq \exp\left(-\frac{\pi^2 N^2}{2 \ln N} (c - 1/2)^2\right)$$

$$\Delta(N) = \left\langle \left(\mathcal{N}_+ - \frac{N}{2} \right)^2 \right\rangle \approx \frac{\ln N}{\pi^2}$$



Tails



$$\mathcal{P}(\mathcal{N}_+ = N, N) \simeq \exp(-\beta N^2 \theta), \quad \theta = (\ln 3)/4$$

MAIN RESULT FOR LARGE N

$$\mathcal{P}(\mathcal{N}_+ = cN, N) \approx \exp(-\beta N^2 \Phi(c))$$

$$\lim_{N \rightarrow \infty} \left[\frac{-\log \mathcal{P}(\mathcal{N}_+ = cN, N)}{\beta N^2} \right] = \Phi(c)$$

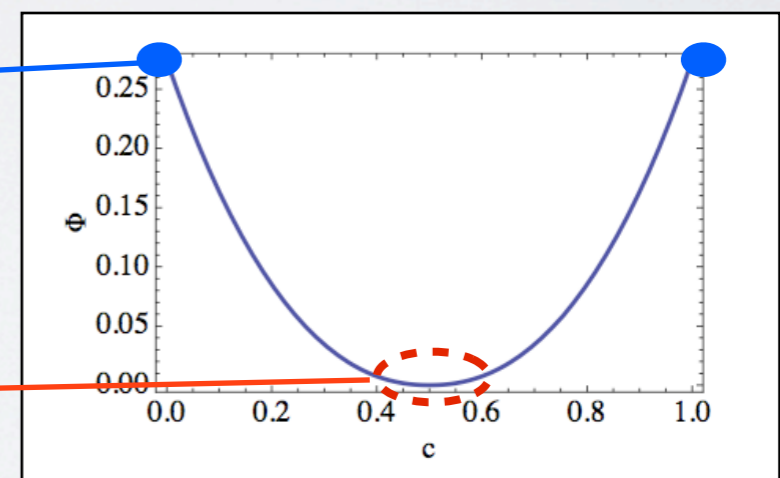
in agreement with
Dean & Majumdar

$$\Phi(0) = \Phi(1) = \theta = \log 3/4 \approx 0.27..$$

$$\Phi(c = 1/2 + \delta) = -\frac{\pi^2}{2} \frac{\delta^2}{\log \delta}$$

$$\mathcal{P}(\mathcal{N}_+, N) \approx \exp \left[-\frac{\beta \pi^2}{2 \ln(N)} (\mathcal{N}_+ - N/2)^2 \right]$$

in agreement with
Cavagna *et al.*

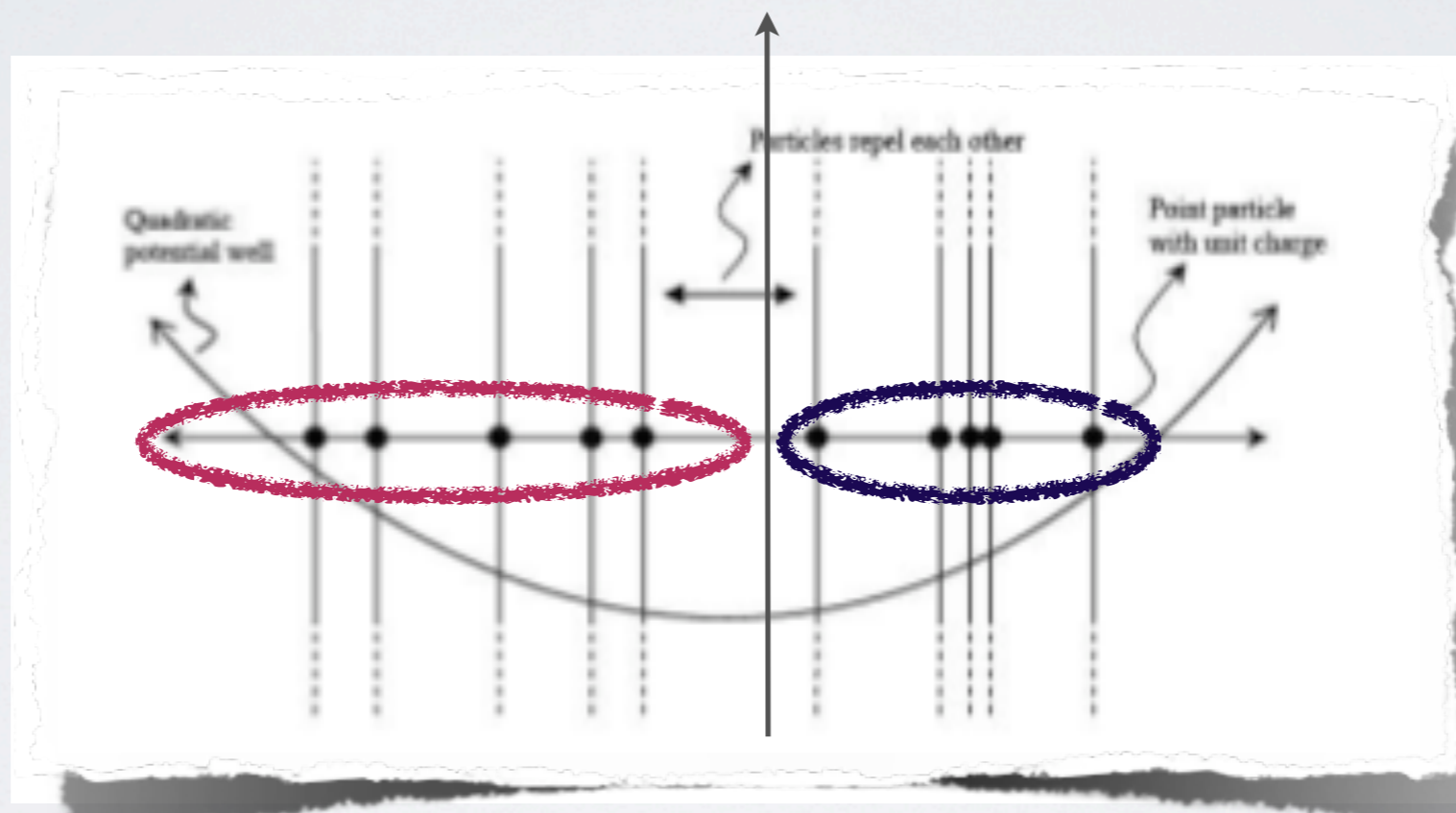


$$\mathcal{P}(\mathcal{N}_+, N) = \int_{-\infty}^{\infty} d\lambda_1 \cdots d\lambda_N P_\beta(\lambda_1, \dots, \lambda_N) \delta \left(\mathcal{N}_+ - \sum_{i=1}^N \theta(\lambda_i) \right)$$

$$e^{-\beta \mathcal{H}(\vec{\lambda})}$$

$$\mathcal{H}(\vec{\lambda}) = \frac{1}{2} \sum_{i=1}^N \lambda_i^2 - \frac{1}{2} \sum_{j \neq k} \log |\lambda_j - \lambda_k|$$

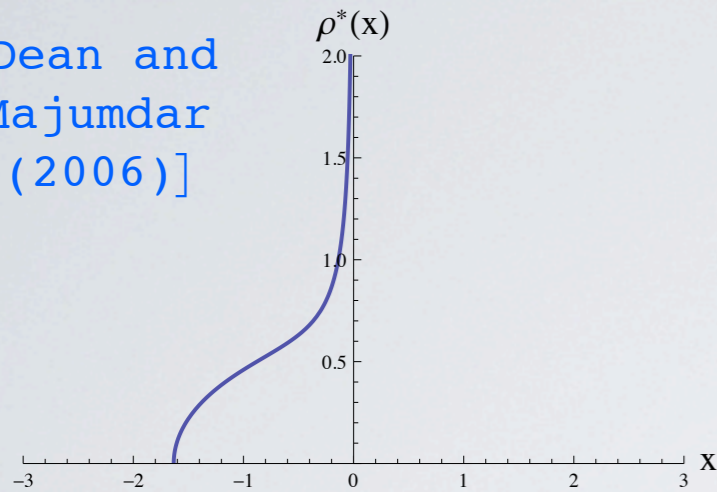
Task: evaluate this integral for large N by mapping it to a Coulomb gas problem



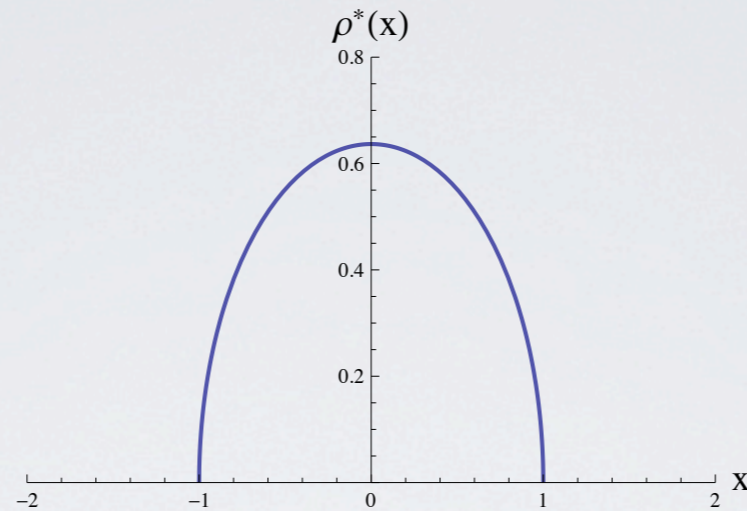
$\mathcal{P}(\mathcal{N}_+, N)$ is the canonical partition function of an auxiliary Coulomb gas with an extra hard constraint

PHASE TRANSITIONS IN THE CONSTRAINED GAS

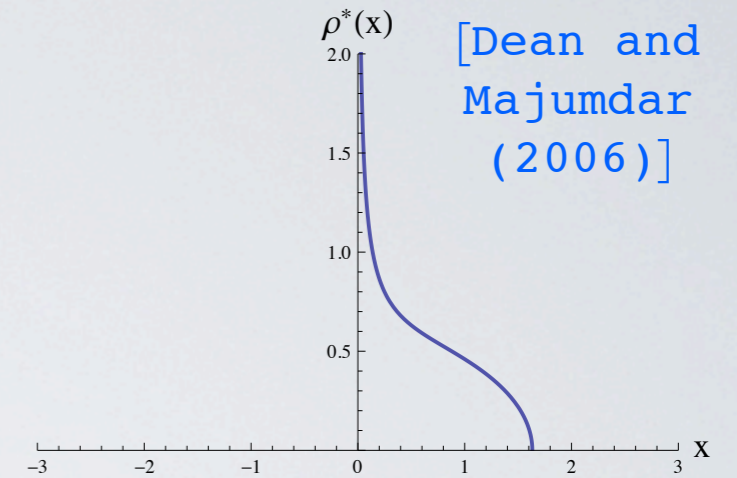
[Dean and Majumdar (2006)]



$c = 0$



$c = \frac{1}{2}$

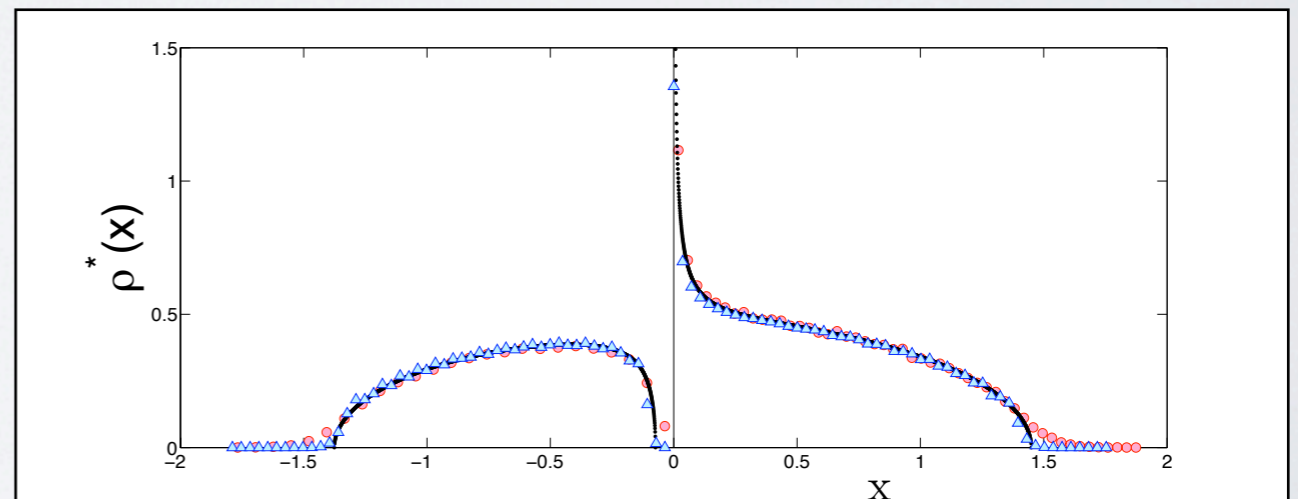


[Dean and Majumdar (2006)]

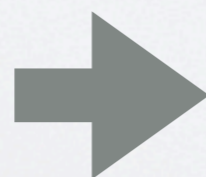
$c = 1$



$$\rho^*(x) = \frac{1}{\pi} \sqrt{\frac{L-x}{x} (x + L/a)(x + (1 - 1/a)L)}$$



$\rho^*(x)$



Partition Function



$\mathcal{P}(\mathcal{N}_+, N)$

$$\mathcal{P}(\mathcal{N}_+, N) = \int_{-\infty}^{\infty} d\lambda_1 \cdots d\lambda_N P_\beta(\lambda_1, \dots, \lambda_N) \delta \left(\mathcal{N}_+ - \sum_{i=1}^N \theta(\lambda_i) \right)$$

$$\mathcal{P}(\mathcal{N}_+, N) \propto \int \mathcal{D}[\rho] e^{-\beta N^2 \mathcal{F}_c[\rho]}$$

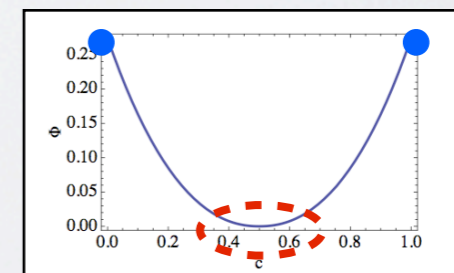
where

$$\begin{aligned} \mathcal{F}_c[\rho] = & \frac{1}{2} \int_{-\infty}^{\infty} dx x^2 \rho(x) - \frac{1}{2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} dx dx' \rho(x) \rho(x') \ln |x - x'| + \\ & + A_1 \left(\int_{-\infty}^{\infty} dx \theta(x) \rho(x) - c \right) + A_2 \left(\int_{-\infty}^{\infty} dx \rho(x) - 1 \right) \end{aligned}$$

Saddle point of the free energy: equilibrium density of the fluid

$$\rho^*(x)$$

$$\mathcal{P}(\mathcal{N}_+ = cN, N) \simeq \exp \left[-\beta N^2 \left(\underbrace{\mathcal{F}_c[\rho^*] - \mathcal{F}_{1/2}[\rho^*]}_{\Phi(c)} \right) \right]$$



SOLVING THE SADDLE-POINT EQUATION

$$\frac{\delta \mathcal{F}_c[\rho]}{\delta \rho} = 0 \quad \Rightarrow \quad \rho^*(x)$$

Inverse
Electrostatic
Problem

$$x^2 + A_1 \theta(x) + A_2 = 2 \int_{-\infty}^{\infty} \rho^*(y) \ln |x - y| dy$$

$$x = \text{Pr} \int dy \frac{\rho^*(y)}{x - y}$$

Equazioni integrali singolari del tipo di Carleman.

FRANCESCO G. TRICOMI (a Torino).

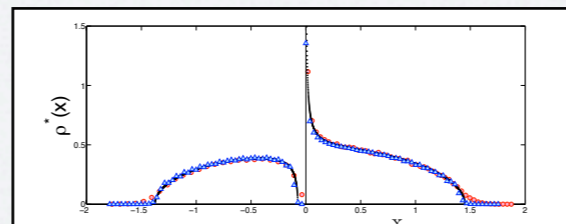
A Mauro Picone nel suo 70^{mo} compleanno.

Commun. math. Phys. 59, 35—51 (1978)

Communications in
Mathematical
Physics
© by Springer-Verlag 1978

Planar Diagrams

E. Brézin, C. Itzykson, G. Parisi*, and J. B. Zuber
Service de Physique Théorique, Centre d'Études Nucléaires de Saclay, F-91190 Gif-sur-Yvette, France



New!
Iterated single-
support solution
by Tricomi (1957)

[Brezin et al. 1978]

Phys. Rev. E **83**, 041105 (2011)

IN SUMMARY...

$$\mathcal{P}(N_+ = cN, N) \simeq \exp \left[-\beta N^2 \left(\underbrace{\mathcal{F}_c[\rho^*] - \mathcal{F}_{1/2}[\rho^*]}_{\Phi(c)} \right) \right]$$

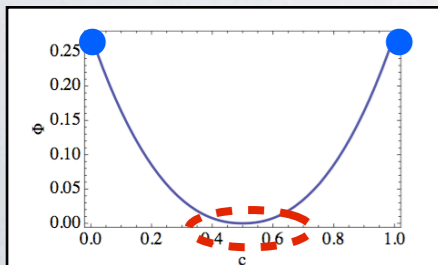
where

$$\rho^*(x) = \frac{1}{\pi} \sqrt{\frac{L-x}{x} (x + L/a)(x + (1 - 1/a)L)}$$

and

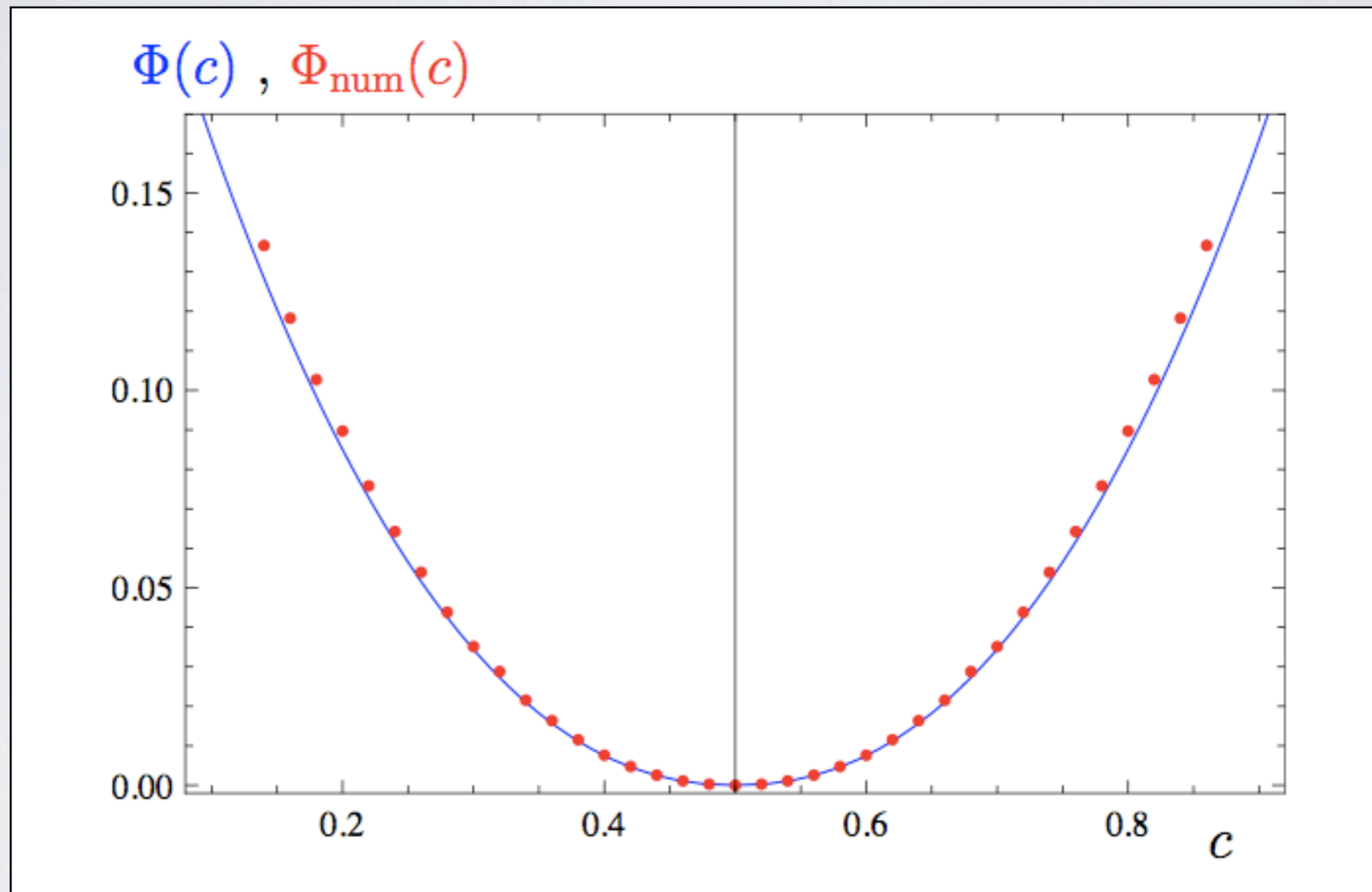
$$\begin{aligned} \mathcal{F}_c[\rho] = & \frac{1}{2} \int_{-\infty}^{\infty} dx x^2 \rho(x) - \frac{1}{2} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} dx dx' \rho(x) \rho(x') \ln |x - x'| + \\ & + A_1 \left(\int_{-\infty}^{\infty} dx \theta(x) \rho(x) - c \right) + A_2 \left(\int_{-\infty}^{\infty} dx \rho(x) - 1 \right) \end{aligned}$$

$$\Phi(c) = \frac{1}{4} [L^2 - 1 - \log(2L^2)] + \frac{(1-c)}{2} \log(a) - \frac{(1-c)(a^2-1)}{4a^2} L^2 + \frac{c}{2} \int_L^{\infty} W_1(x) dx + \frac{(1-c)}{2} \int_{L/a}^{\infty} W_2(x) dx$$



$$W_1(x) = F(x) - \frac{1}{x} = x - \frac{1}{x} - \sqrt{\frac{(x-L)}{x} \left(x + \frac{L}{a} \right) \left(x + \left(1 - \frac{1}{a} \right) L \right)}$$

Some numerics...



SUMMARY

- Any matrix coming up in your research?
Randomize! (and come to my office later...)
- Strongly correlated random variables
- Ubiquity - Universality of local statistics
- Rare events for strongly correlated random variables ---> exactly solvable cases!

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- S.N. Majumdar and PV, [arXiv: 1112.5391] (2011)
- A. Cavagna, J.P. Garrahan, and I. Giardinà, Phys. Rev. B **61**, 3960 (2000)
- D. Dean and S.N. Majumdar, Phys. Rev. Lett. **97**, 160201 (2006); Phys. Rev. E **77**, 041108 (2008)
- N.S. Witte and P.J. Forrester, [arXiv: 1110.0119] (2011)

Thank you.